
Advances in Models for Multivariate Nominal or Ordinal Variables with Latent Variables

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Models for Nominal & Ordinal Variables

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● Models for Nominal & Ordinal Variables

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- Example data set: Espelage et al.
- Review of existing approaches for analyzing such data.
- Log-multiplicative association models.
 - ◆ Underlying models that lead to LMA
 - ◆ Conditional Specification.
- The State of the Art:
 - ◆ Multiple correlated latent variables.
 - ◆ Restrictions on scale values for response options and location parameters.
 - ◆ Covariates.
 - ◆ Estimation.
- Areas for future work.
- Time permitting, a nominal example.



The Espelage, Holt & Henkel Data (2004)

Data from Espelage, D.L., Holt, M.K., & Henkel, R.R. (2004). Examination of peer-group contextual effects on aggression during early adolescence. *Child Development*, 74, 205–220.

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The data consist of responses of students from a midwestern middle school to items of the Illinois Victimization Scale (Espelage & Holt, 2001).

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Bully Sub-scale Items:

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Fight Sub-scale Items:

- I got in a physical fight.
- I threatened to hurt or hit another student.
- I hit back when someone hits me first.

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Response scale:

Never, 1 or 2 times, 3 or 4 times, 5 or 6 times, and 7 or more times.

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The role Gender? Do girls tend to be more verbal bullies and boys more physical? ... *The findings are mixed...*

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Review of Major Existing Approaches

to Latent Variable Modeling of Discrete Response Data:

- Quantify the data and then use factor analysis (or SEM) for continuous data.

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 - ◆ Multiple latent variables is a problem for standard estimation algorithms (i.e., numerical integration).
- Factor analysis of discrete data (Bartholomew, Steele, Moustaki & Galbraith, 2008)
 - ◆ Lack of available of software and flexibility of implementation.
 - ◆ Methods and programs for nominal data are sorely lacking and “...work on ordinal categorical variables is nearer the research frontier and is consequently more incomplete, and in some sense, more difficult than other methods.” (p. 243)



Log-multiplicative Association Models

- Structured Poisson regression model with 2-way interactions.

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Log-multiplicative Association Models

- Structured Poisson regression model with 2-way interactions.
- Generalization of Goodman's (1979, 1986) $RC(M)$ association model for two-way tables to multi-way tables, i.e.,

$$\log(P(y_i = j_i, y_k = \ell_k)) = \lambda + \lambda_{ij_i}^R + \lambda_{kl_k}^C + \sum_m \phi_m \nu_{ij_i m}^R \nu_{kl_k m}^C$$

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- When equally spaced scores are input for the ν 's (and $M = 1$), then the model is known as a the **uniform association model**.

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- Takane (1987): **Ideal point discriminant analysis** without a centroid restriction on the columns (criterion groups) is equivalent to the *RC* association model.

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- Andersen (1995): **Rasch model for polytomous items** where an item's response options are the rows and the columns are (categorical) estimates of ability/latent trait.
- The **general log-multiplicative association (LMA)** model

$$\log(P(\mathbf{y})) = \lambda + \sum_i \lambda_{ij} + \sum_i \sum_{k>i} \sum_m \sum_{m' \geq m} \sigma_{mm'} \nu_{ijm} \nu_{klm'}$$

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Underlying Models that Imply a LMA Model

that I know of...

- Are implied by underlying [multivariate normal distribution](#) (Goodman, 1979; Becker, 1989; others).

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- Can be derived from a **formative latent variable model** using statistical graphical models (Anderson & Böckenholt, 2000; Anderson & Vermunt, 2000; Anderson, 2002; Anderson & Tettaah, 2007).

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- Can be derived from a **formative latent variable model** using statistical graphical models (Anderson & Böckenholt, 2000; Anderson & Vermunt, 2000; Anderson, 2002; Anderson & Tettaah, 2007).
- Can be derived from a **reflective latent variable model** using standard item response theory methodology (Anderson & Yu, 2007; Anderson, Li, & Vermunt, 2007; Anderson, Verkuilen & Peyton, in press).

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Got in fight

Threatened to hurt

Hit back

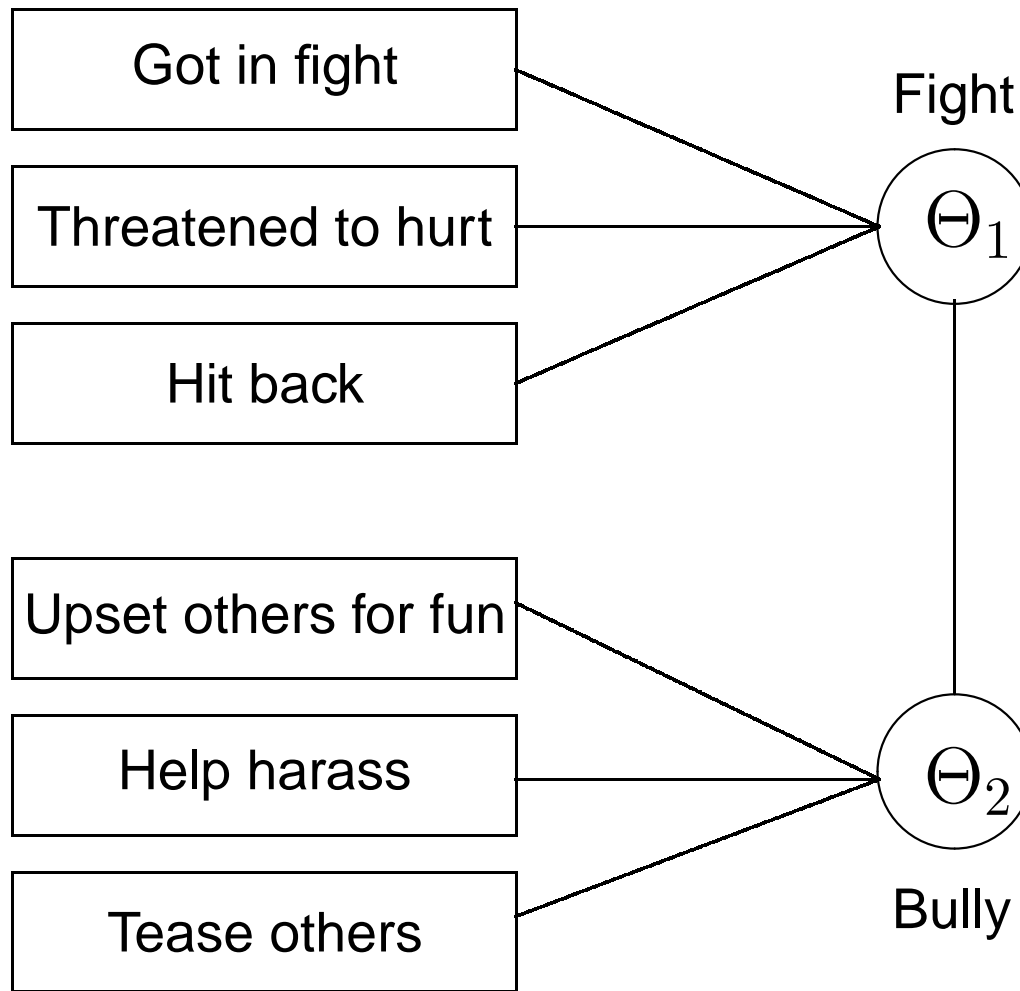
Upset others for fun

Help harass

Tease others



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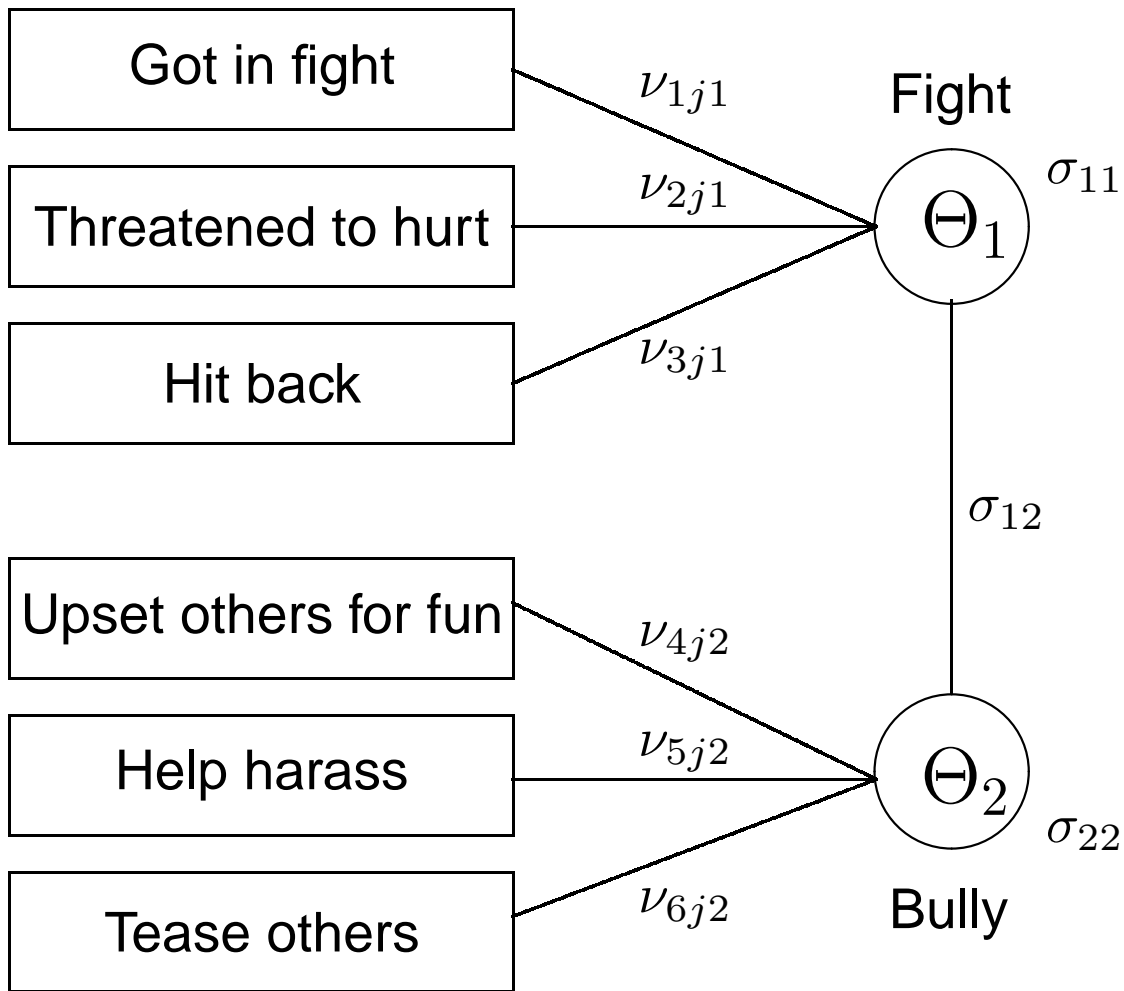
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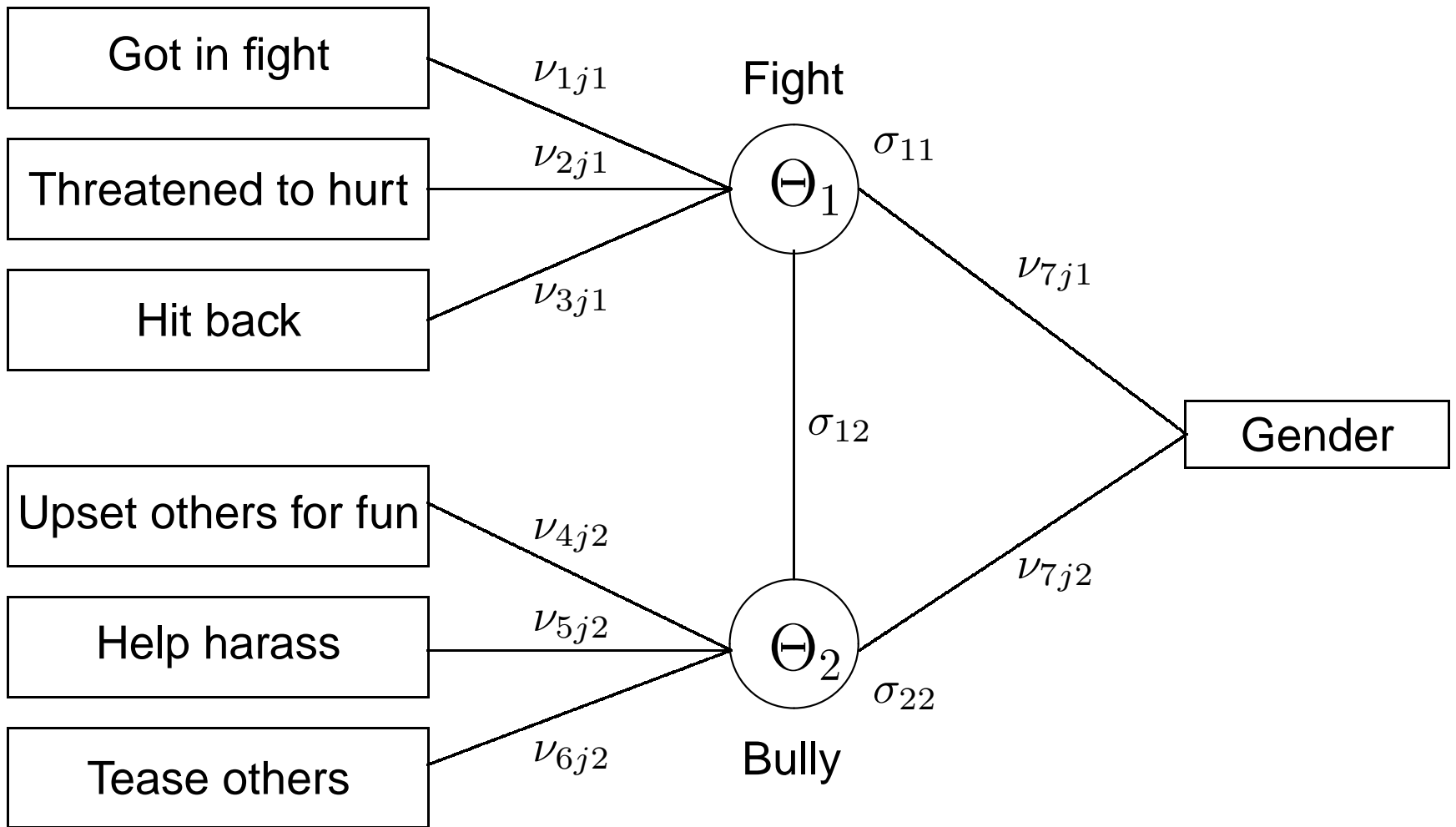
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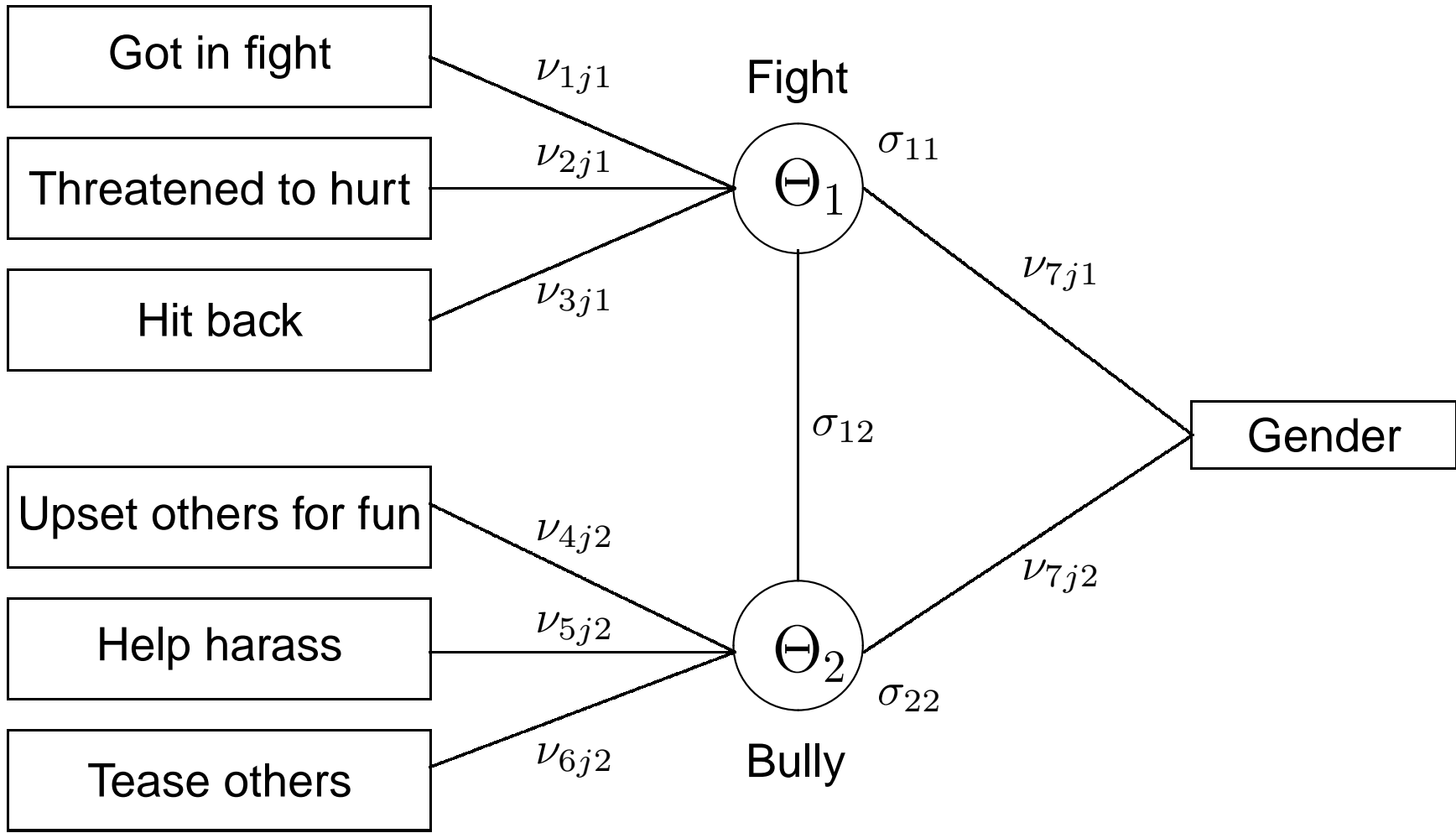


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Assumptions & Implications

- The response pattern y follows a multinomial distribution.

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Assumptions & Implications

- The response pattern y follows a multinomial distribution.
- Absence of a line connecting variables indicates conditional independence

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Assumptions & Implications

- The response pattern \mathbf{y} follows a multinomial distribution.
- Absence of a line connecting variables indicates conditional independence
- For all possible response patterns \mathbf{y} ,

$$\Theta | \mathbf{y} \sim MVN(\boldsymbol{\mu}_{\mathbf{y}}, \boldsymbol{\Sigma}) \text{ i.i.d.}$$

where

$$\boldsymbol{\mu}_{\mathbf{y}} = \begin{pmatrix} \sigma_{11} & \sigma_{12} & \dots & \sigma_{1M} \\ \sigma_{12} & \sigma_{22} & \dots & \sigma_{2M} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{1M} & \sigma_{2M} & \dots & \sigma_{MM} \end{pmatrix} \begin{pmatrix} \sum_i \nu_{ij1} \\ \sum_i \nu_{ij2} \\ \vdots \\ \sum_i \nu_{ijM} \end{pmatrix}$$

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Our example,

$$\text{Fight: } \mu_1 | \mathbf{y} = \sigma_{11} \left(\sum_i \nu_{ij1} \right) + \sigma_{12} \left(\sum_i \nu_{ij2} \right)$$

$$\text{Bully: } \mu_2 | \mathbf{y} = \sigma_{22} \left(\sum_i \nu_{ij2} \right) + \sigma_{12} \left(\sum_i \nu_{ij1} \right)$$

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Conditional Approach

Consider the following conditional logistic regression model,

$$P(Y_i = j | y_{kl}, k \neq i, \mathbf{x}) = \frac{\exp(\lambda_{ij} + \sum_p \beta_{ijp} x_p + \sum_{k \neq i} \psi_{ij|kl})}{\sum_h \exp(\lambda_{ih} + \sum_p \beta_{ihp} x_p + \sum_{k \neq i} \psi_{ih|kl})}$$

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● Simplifying the Model

● Special Case #1:

$M = 1$

● Special Case #2: M

● Recent Developments: LMA as IRT Models

● Common and Novel IRT Models as LMAs

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where

- λ_{ij} is an intercept or location parameter.
- $\psi_{ij|kl}$ is the parameter for variable y_k when predicting variable y_i .

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where

- λ_{ij} is an intercept or location parameter.
- $\psi_{ij|kl}$ is the parameter for variable y_k when predicting variable y_i .

If we have this model for each item i ($i = 1, \dots, I$) and $\psi_{ij|kl} = \psi_{kl|ij}$, then model for the joint distribution of all items is

$$\log(P(y_{1j}, \dots, y_{Ij} | \mathbf{x})) = \lambda + \sum_i \lambda_{ij} + \sum_i \sum_p \beta_{ijp} x_p + \sum_i \sum_{k > i} \psi_{ij|kl}$$

This is basically a log-linear model with all 2-way interactions.

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Conditional Approach

Consider the following conditional logistic regression model,

$$P(Y_i = j | y_{kl}, k \neq i, \mathbf{x}) = \frac{\exp(\lambda_{ij} + \sum_p \beta_{ijp} x_p + \sum_{k \neq i} \psi_{ij|kl})}{\sum_h \exp(\lambda_{ih} + \sum_p \beta_{ihp} x_p + \sum_{k \neq i} \psi_{ih|kl})}$$

where

- λ_{ij} is an intercept or location parameter.
- $\psi_{ij|kl}$ is the parameter for variable y_k when predicting variable y_i .

If we have this model for each item i ($i = 1, \dots, I$) and $\psi_{ij|kl} = \psi_{kl|ij}$, then model for the joint distribution of all items is

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This is basically a log-linear model with all 2-way interactions.

Proof for the dichotomous, Joe & Liu (1996); for other cases, Anderson, Li & Vermunt (2007), and Anderson, Verkuilen & Peyton (in press).

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Simplifying the Model

For each pair of items, there is a $(J \times L)$ matrix of ψ 's,

$$\Psi_{i|k} = \begin{pmatrix} \psi_{i1|k1} & \psi_{i1|k2} & \cdots & \psi_{i1|kL} \\ \psi_{i2|k1} & \psi_{i2|k2} & \cdots & \psi_{i2|kL} \\ \vdots & \vdots & \ddots & \vdots \\ \psi_{iJ|k1} & \psi_{iJ|k2} & \cdots & \psi_{iJ|kL} \end{pmatrix}$$

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The model can be simplified by considering lower rank decompositions:

$$\Psi_{i|k} = N_i^{[ik]} \Sigma^{[ik]} N_k^{[ik]'} \quad \text{where } \Sigma^{[ik]} \text{ is diagonal.}$$

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Simplifying the Model

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The model can be simplified by considering lower rank decompositions:

$$\Psi_{i|k} = \mathbf{N}_i^{[ik]} \Sigma^{[ik]} \mathbf{N}_k'^{[ik]} \quad \text{where } \Sigma^{[ik]} \text{ is diagonal.}$$

However, here we'll mostly consider those of the form $\Psi_{i|k} = \mathbf{N}_i \Sigma \mathbf{N}_k'$ where Σ is not necessarily diagonal and

$$\underbrace{\mathbf{N}_i}_{(J \times M)} = \begin{pmatrix} \nu_{i11} & \nu_{i12} & \dots & \nu_{i1M} \\ \nu_{i21} & \nu_{i22} & \dots & \nu_{i2M} \\ \vdots & \vdots & \ddots & \vdots \\ \nu_{iJ1} & \nu_{iJ2} & \dots & \nu_{iJM} \end{pmatrix}$$

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Special Case #1: $M = 1$

$$\Psi_{i|k} = \nu_{i1}\sigma_{11}\nu'_{k1} = \{\sigma_{11}\nu_{ij1}\nu_{kl1}\}$$

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Special Case #1: $M = 1$

$$\Psi_{i|k} = \nu_{i1}\sigma_{11}\nu'_{k1} = \{\sigma_{11}\nu_{ij1}\nu_{kl1}\}$$

The conditional logistic regression model for each item i is

$$P(Y_i = j | y_{kl}, k \neq i) = \frac{\exp(\lambda_{ij} + \nu_{ij1}(\sigma_{11} \sum_{k \neq i} \nu_{kl1}))}{\sum_h \exp(\lambda_{ih} + \nu_{ih1}(\sigma_{11} \sum_{k \neq i} \nu_{kl1}))}$$

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Special Case #1: $M = 1$

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The conditional logistic regression model for each item i is

$$\begin{aligned} P(Y_i = j | y_{kl}, k \neq i) &= \frac{\exp(\lambda_{ij} + \nu_{ij1}(\sigma_{11} \sum_{k \neq i} \nu_{kl1}))}{\sum_h \exp(\lambda_{ih} + \nu_{ih1}(\sigma_{11} \sum_{k \neq i} \nu_{kl1}))} \\ &= \frac{\exp(b_{ij} + a_{ij}\tilde{\theta})}{\sum_h \exp(b_{ih} + a_{ih}\tilde{\theta})} \end{aligned}$$

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- $\lambda_{ij} = b_{ij}$ is an intercept or “difficulty” parameter.

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$$\Psi_{i|k} = \nu_{i1}\sigma_{11}\nu'_{k1} = \{\sigma_{11}\nu_{ij1}\nu_{kl1}\}$$

The conditional logistic regression model for each item i is

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- $\lambda_{ij} = b_{ij}$ is an intercept or “**difficulty**” parameter.
- $\nu_{ij1} = a_{ij}$ is a slope or “**discrimination**” parameter.



Special Case #1: $M = 1$

$$\Psi_{i|k} = \nu_{i1}\sigma_{11}\nu'_{k1} = \{\sigma_{11}\nu_{ij1}\nu_{kl1}\}$$

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- **Bock’s nominal response model and all its special cases.**

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- $\lambda_{ij} = b_{ij}$ is an intercept or “**difficulty**” parameter.
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■ **Bock’s nominal response model and all its special cases.**

■ The LMA

$$P(\mathbf{y}) = \lambda + \sum_i \lambda_{ij} + \sigma_{11} \sum_i \sum_{k > i} \nu_{ij1} \nu_{kl1}$$



Special Case #2: M

$$\Psi_{i|k} = N_i \Sigma N'_k = \left\{ \sum_m \sum_{m'} \nu_{ijm} \sigma_{mm'} \nu_{klm'} \right\}$$

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Special Case #2: M

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$$= \frac{\exp(b_{ij} + \sum_m a_{ijm} \tilde{\theta}_m)}{\sum_h \exp(b_{ih} + \sum_m a_{ihm} \tilde{\theta}_m)}$$



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$$= \frac{\exp(b_{ij} + \sum_m a_{ijm} \tilde{\theta}_m)}{\sum_h \exp(b_{ih} + \sum_m a_{ihm} \tilde{\theta}_m)}$$

■ $\nu_{ijm} = a_{ijm}$ is the slope or discrimination parameter for variable m .



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- $\nu_{ijm} = a_{ijm}$ is the slope or discrimination parameter for variable m .
- $\tilde{\theta}_m$ is predictor variable m or weighted sum of **rest-scores** or **test-totals**:

$$\tilde{\theta}_m = \underbrace{\sigma_{mm} \left(\sum_{k \neq i} \nu_{klm} \right)}_{\text{rest-score } m} + \sum_{m' \neq m} \sigma_{mm'} \underbrace{\left(\sum_{k \neq i} \nu_{klm'} \right)}_{\text{rest-score or test-total } m'}$$



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- **Multidimensional compensatory IRT model for polytomous items.**



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■ Anderson, Li & Vermunt (2007):

- ◆ Models in the Rasch family—dichotomous and polytomous items, uni- and multi-dimensional latent variables.
- ◆ Pseudo-likelihood estimation.



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- Anderson, Li & Vermunt (2007):
 - ◆ Models in the Rasch family—dichotomous and polytomous items, uni- and multi-dimensional latent variables.
 - ◆ Pseudo-likelihood estimation.

- Anderson & Yu (2007):
 - ◆ Dichotomous items, 1 underlying latent variable.
 - ◆ Different underlying alternative marginal distributions of the latent variable.



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■ Anderson, Li & Vermunt (2007):

- ◆ Models in the Rasch family—dichotomous and polytomous items, uni- and multi-dimensional latent variables.
- ◆ Pseudo-likelihood estimation.

■ Anderson & Yu (2007):

- ◆ Dichotomous items, 1 underlying latent variable.
- ◆ Different underlying alternative marginal distributions of the latent variable.

■ Anderson, Verkuilen & Peyton (in press):

- ◆ Multicategory items and two latent variables (also 3 latent variable and higher order models, but these aren't in the paper).
- ◆ Covariate that influenced the choice of the “don't know” response option (i.e., instructions). The covariate came in from the “left”.
- ◆ Hybrid model.
- ◆ Equality restrictions on some parameters.
- ◆ Models fit using SAS/NLP.



Common and Novel IRT Models as LMAs

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- Special Case #2: M
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Restrictions on LMA Parameters

Common models	$\lambda_{ij} (b_{ij})$	$\nu_{ijm} (a_{ijm})$
2PL	none	none
Nominal response model	none	none
Multidimensional compensatory	none	none
Rasch family	none	input/fixed (ordered)
Graded \times Nominal response	none	ordinal



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Restrictions on LMA Parameters

Common models

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$\nu_{ijm} (a_{ijm})$

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none

Nominal response model

none

none

Multidimensional compensatory

none

none

Rasch family

none

input/fixed (ordered)

Graded \times Nominal response

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Novel ones can be created by modeling or placing restrictions on location parameters (i.e., λ_{ij}), category scores (i.e., ν_{ijm}), and/or σ_{mm} s:



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- Input/fixed.
- Equality.
- Ordinal.
- Linear functions (e.g., $\nu_{ij} = \omega_i x_{ij}$ where $x_{ij} = 0, 1, \dots, J$ or any values).



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- Linear functions (e.g., $\nu_{ij} = \omega_i x_{ij}$ where $x_{ij} = 0, 1, \dots, J$ or any values).
- Set minimum and maximum (e.g., 0 and 1) and estimate scores in between.
- Model σ_{mm} (e.g., $\sigma_{mm} = \sigma_{mm}^* + \beta_m x$ or $\sigma_{mm} = \sigma_{mm}^* \beta_m x$).



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Novel ones can be created by modeling or placing restrictions on location parameters (i.e., λ_{ij}), category scores (i.e., ν_{ijm}), and/or σ_{mm} s:

- Input/fixed.
- Equality.
- Ordinal.
- Linear functions (e.g., $\nu_{ij} = \omega_i x_{ij}$ where $x_{ij} = 0, 1, \dots, J$ or any values).
- Other: e.g., $\nu_{ij} = \omega_i \nu_{ij}^*$ where $\sum_j \nu_{ij}^{*2} = 1$ or set min and max ν_{ij}^* .
- Set minimum and maximum (e.g., 0 and 1) and estimate scores in between.
- Model σ_{mm} (e.g., $\sigma_{mm} = \sigma_{mm}^* + \beta_m x$ or $\sigma_{mm} = \sigma_{mm}^* \beta_m x$).



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● -2Inlike versus Number Parameters

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● -2Inlike versus Number Parameters w/ Gender

● Estimated Item Scale Values

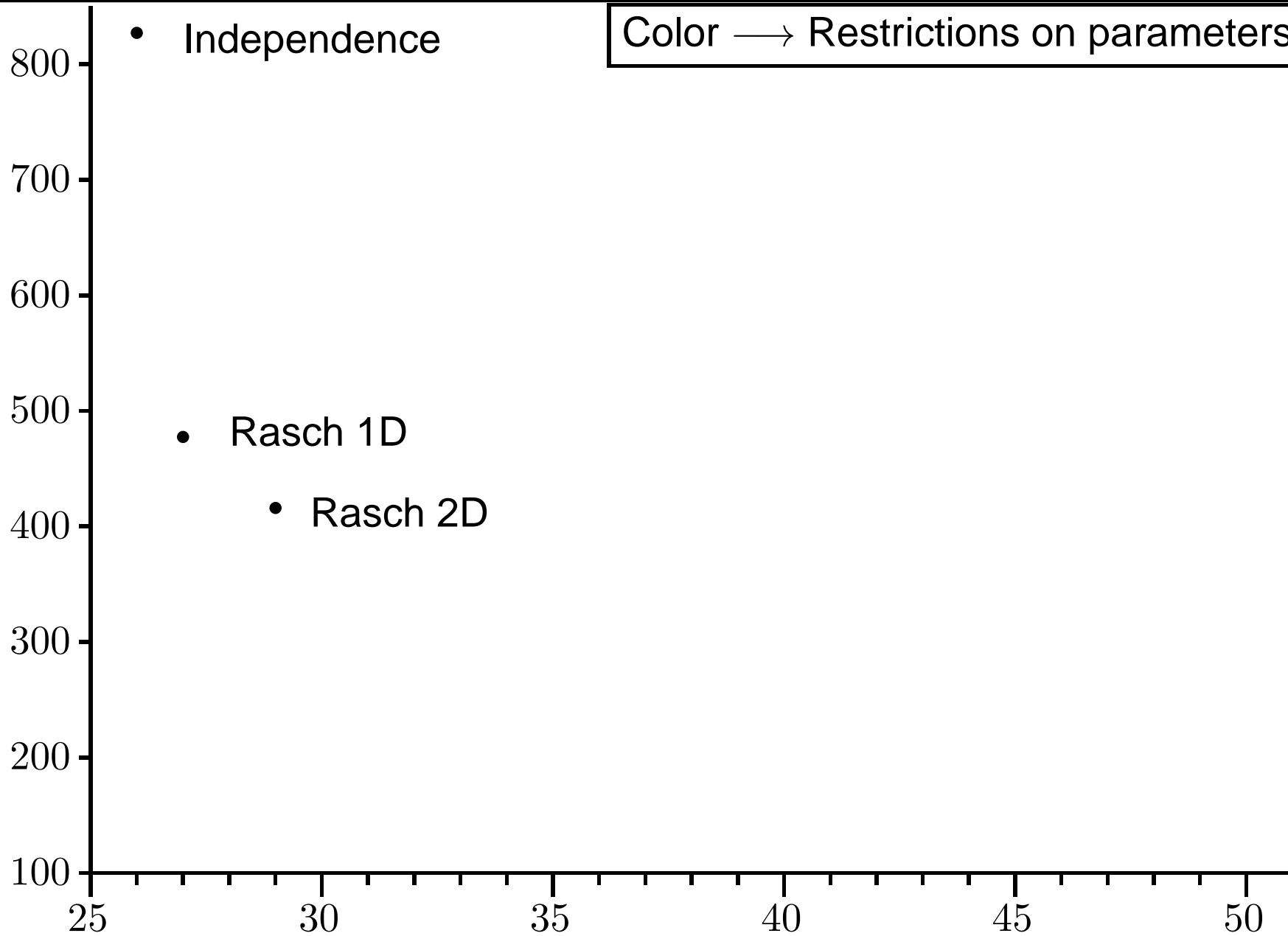
● Item Characteristic Curves

● Item Cumulative Probability Curves

● Estimated Scale Values for Gender

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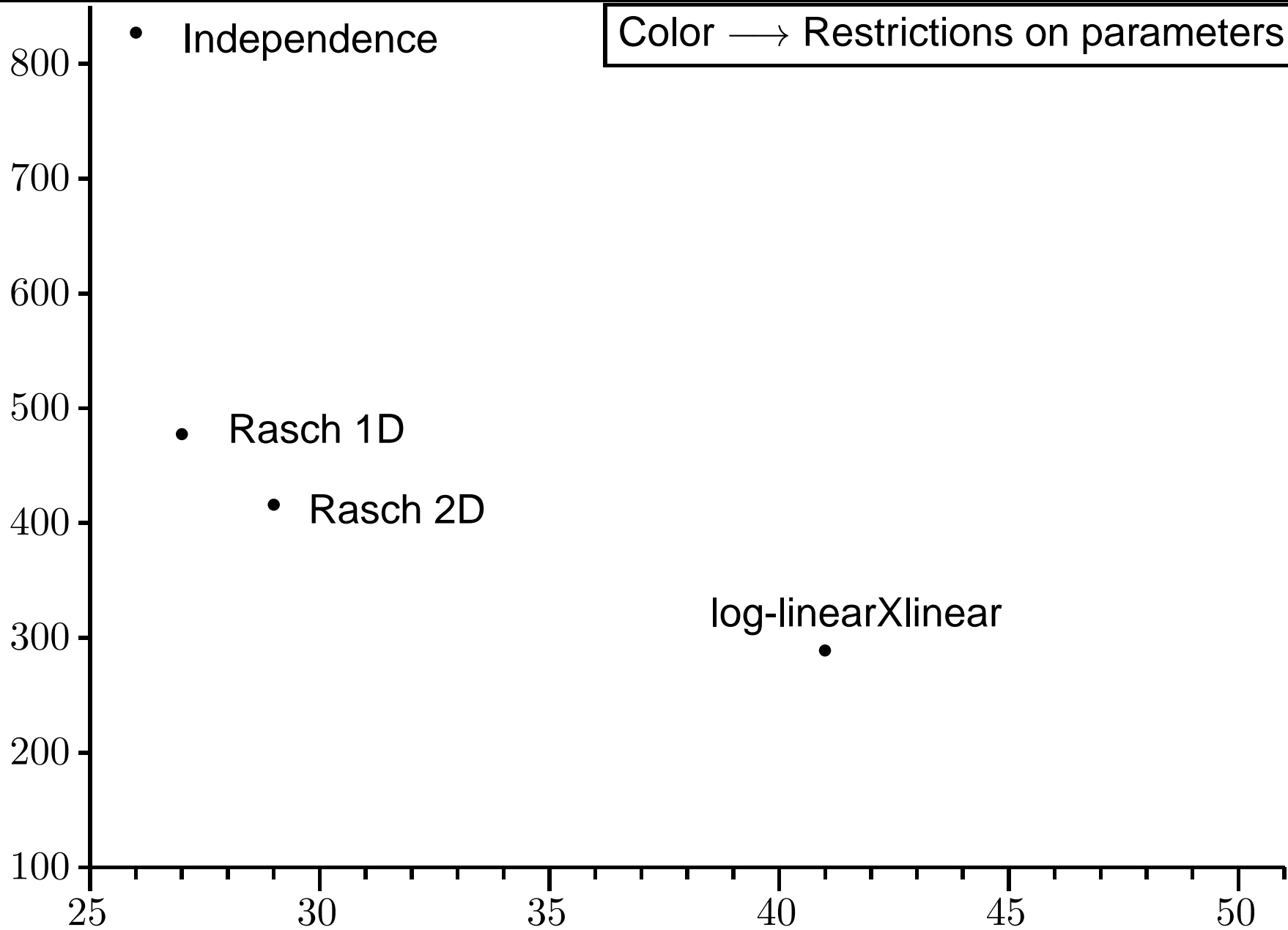




-2Inlike versus Number Parameters

Color → Restrictions on parameters

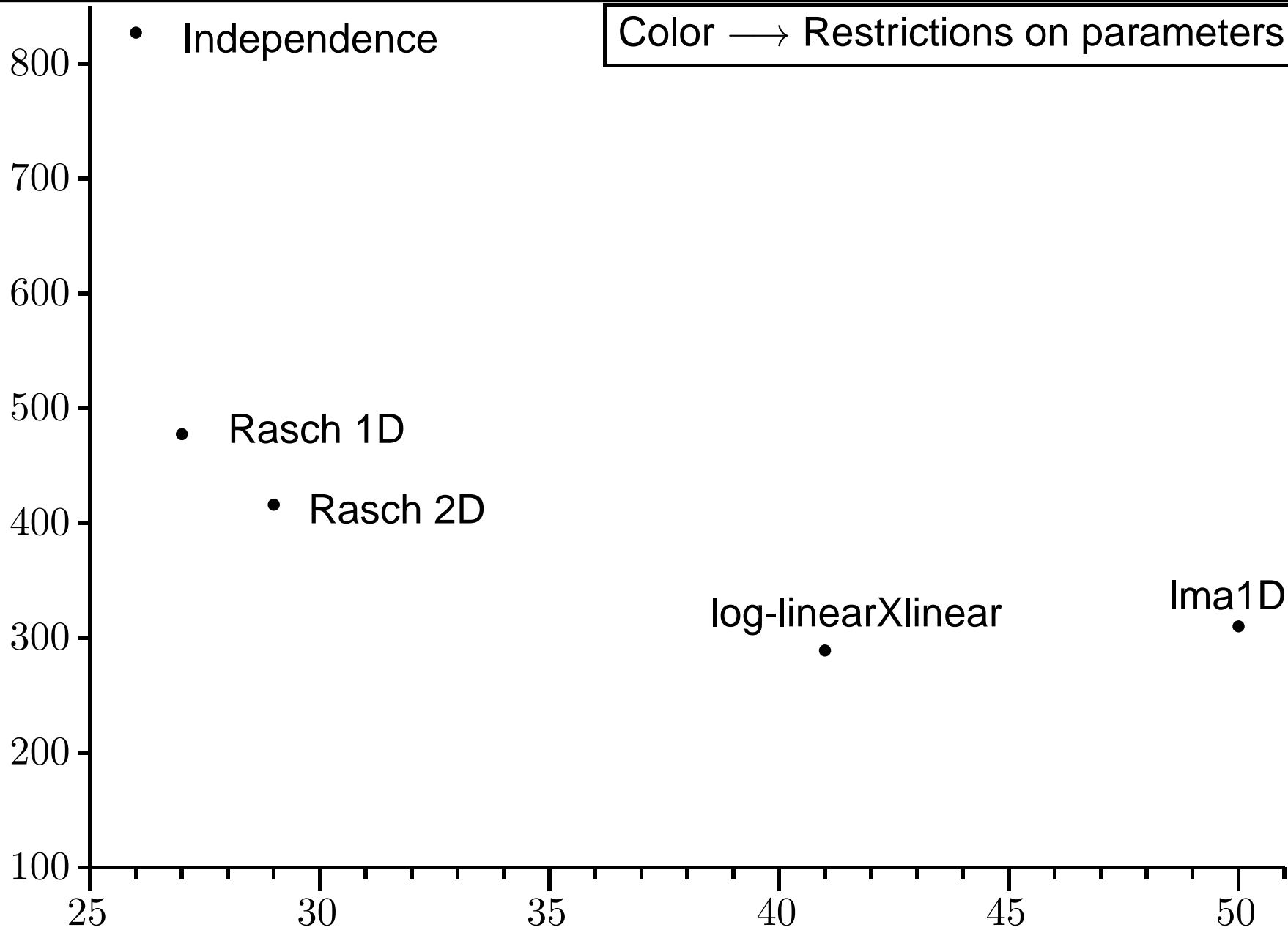
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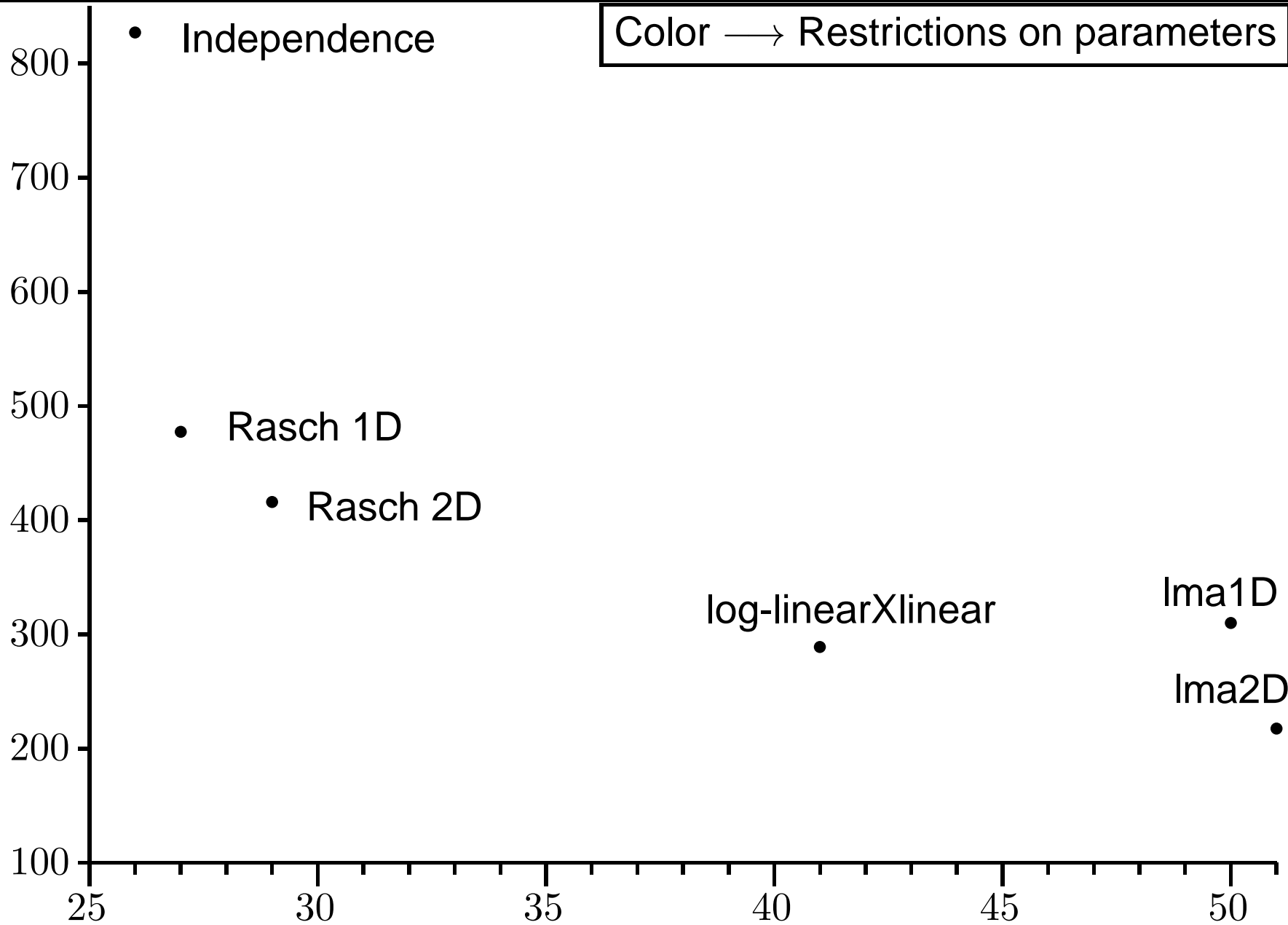
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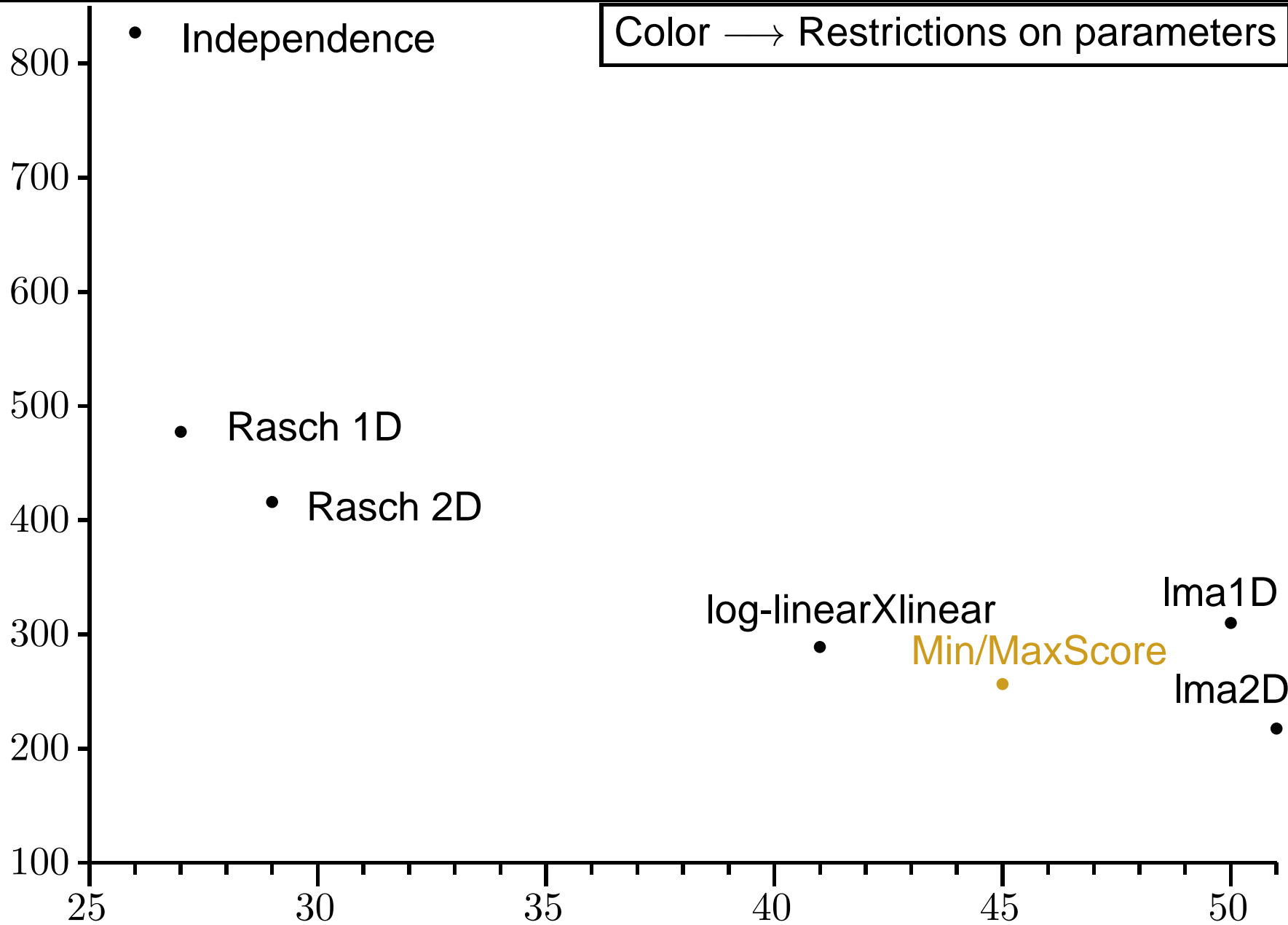
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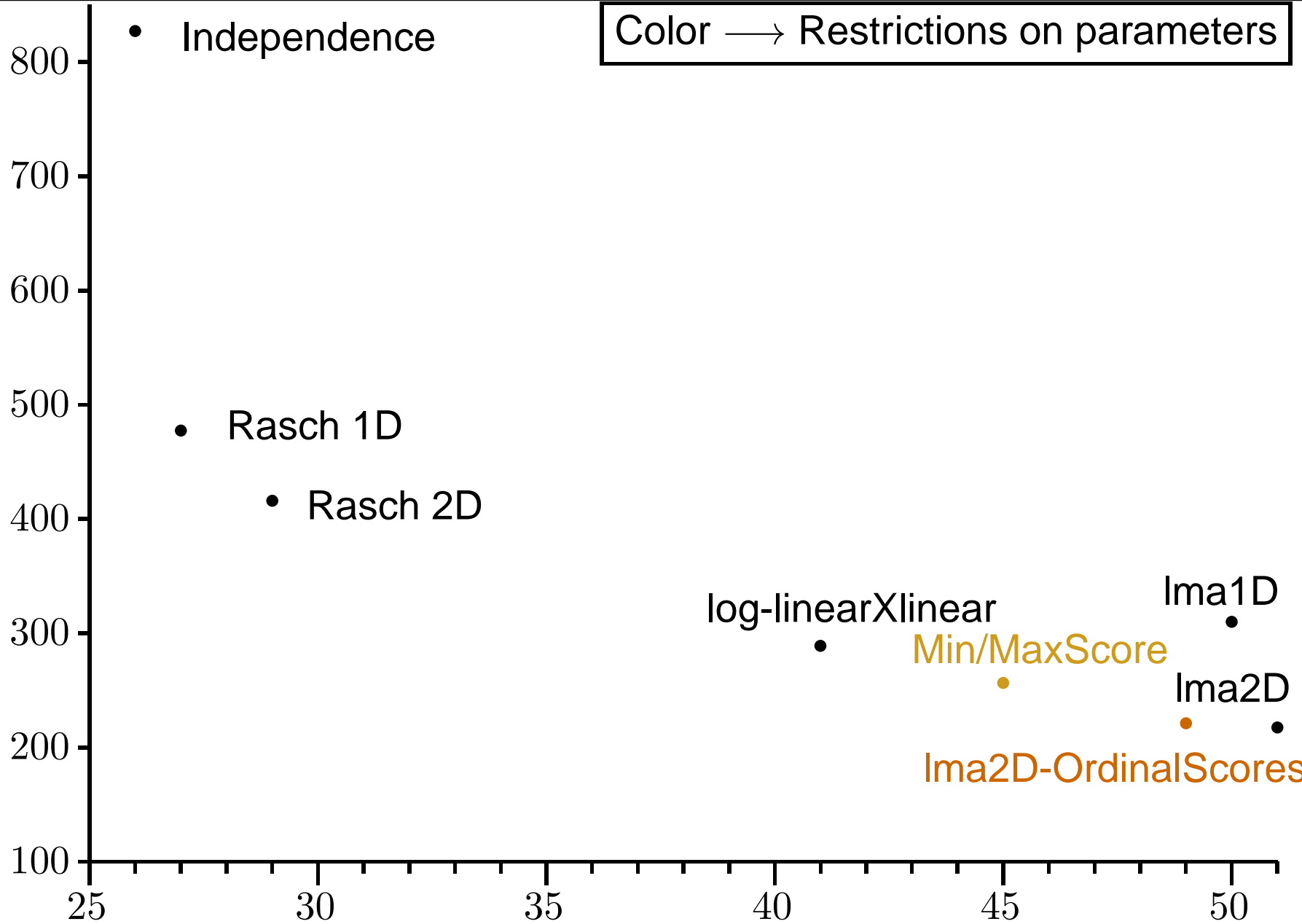




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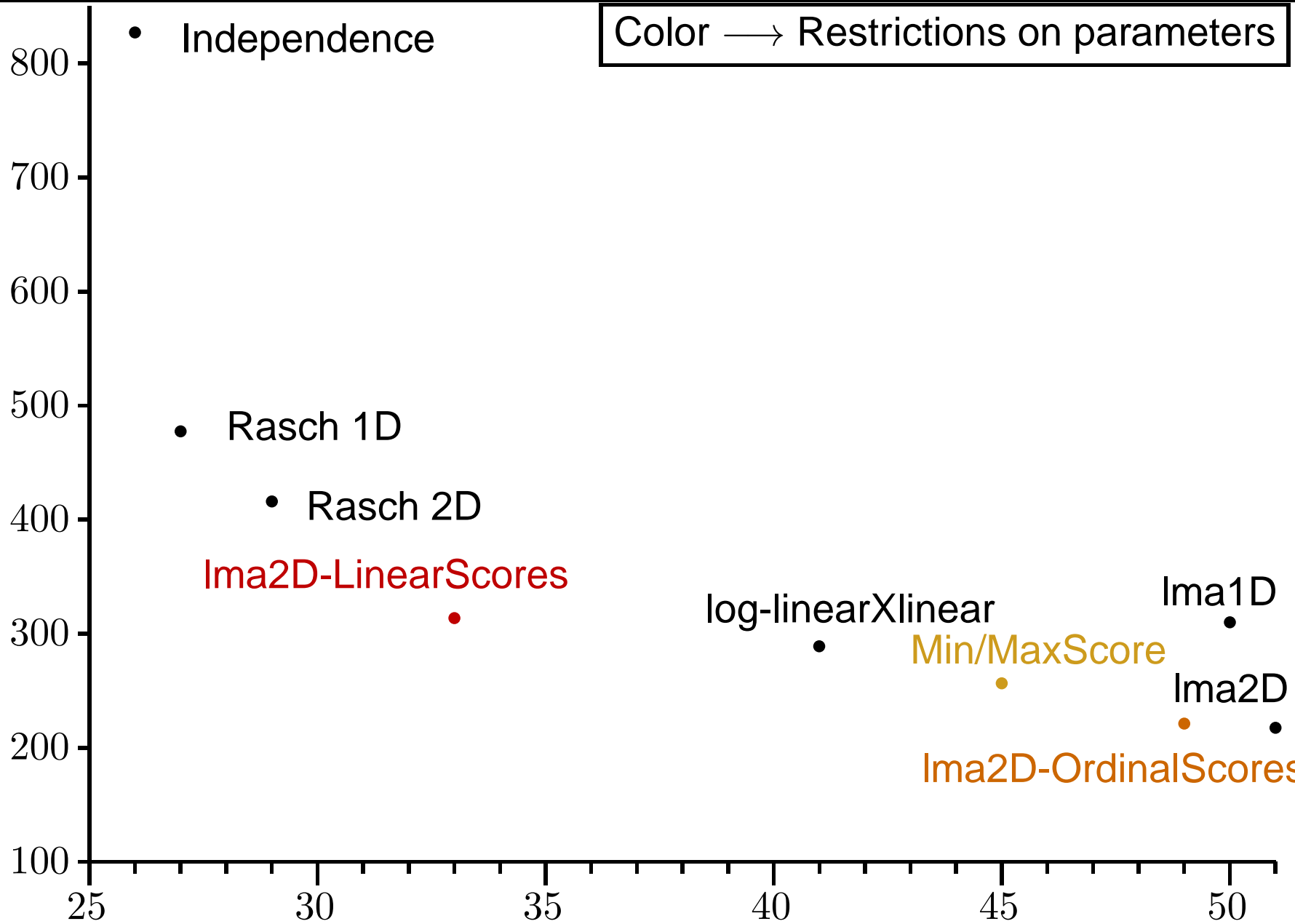
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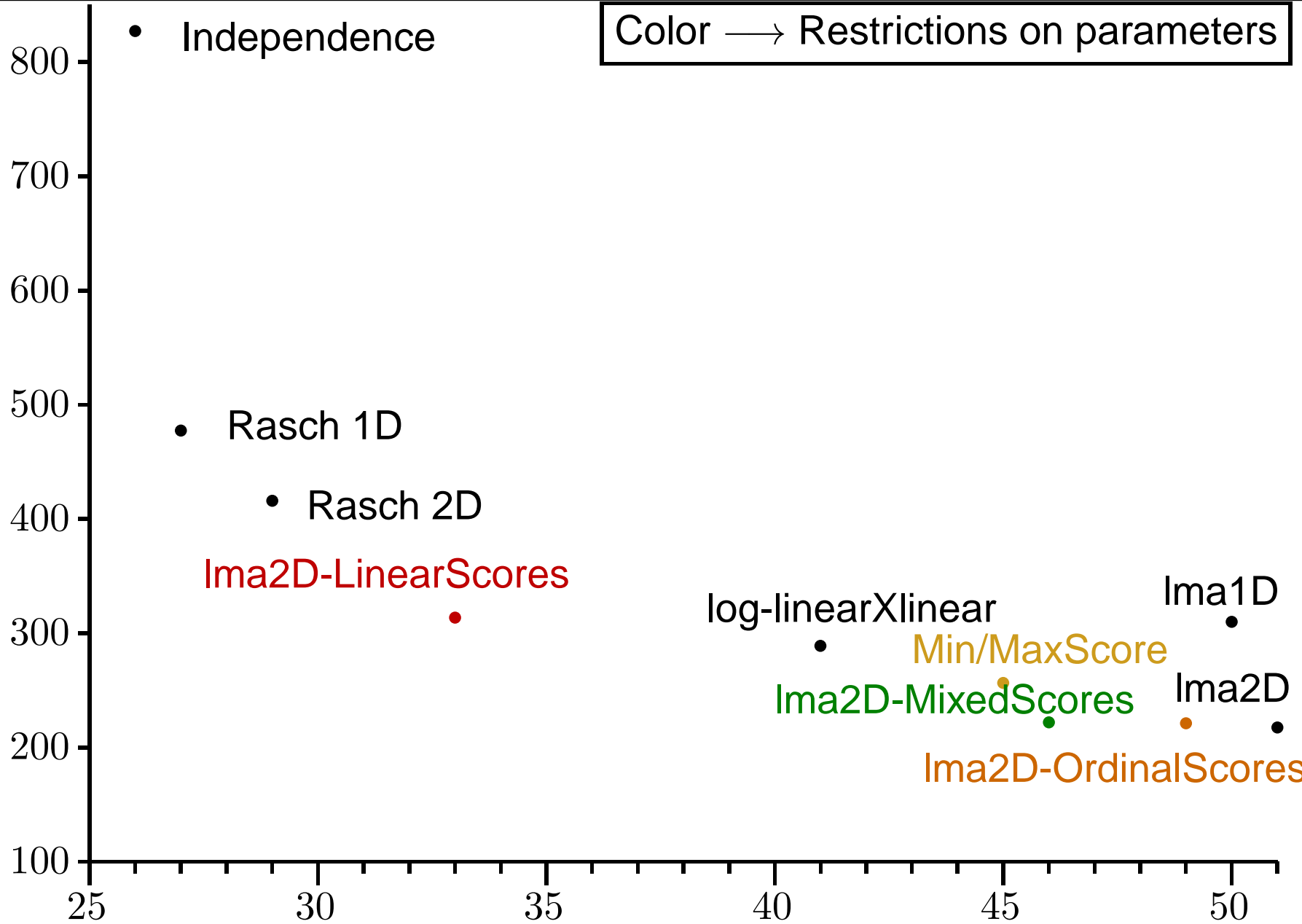




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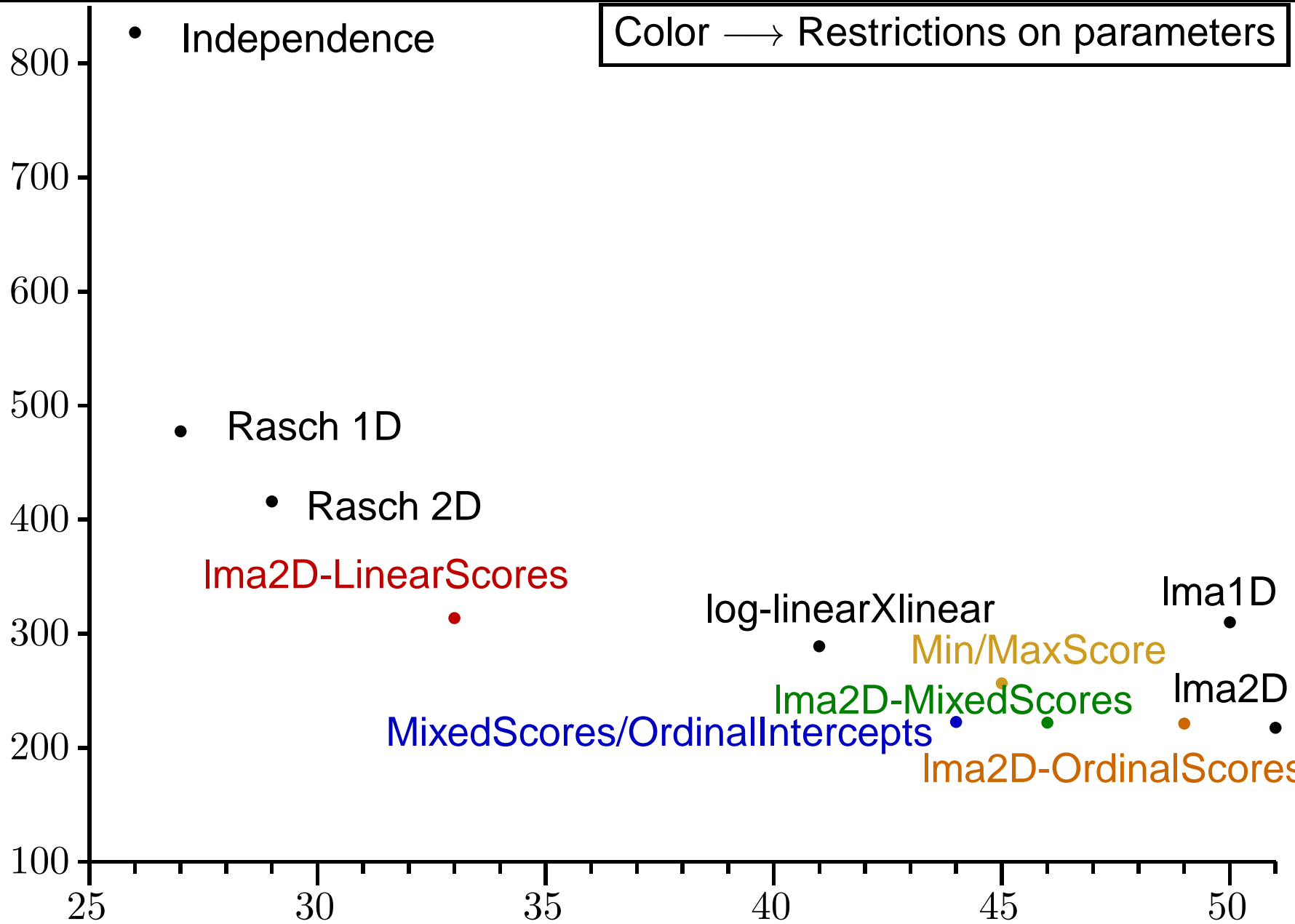
Color → Restrictions on parameters





-2Inlike versus Number Parameters

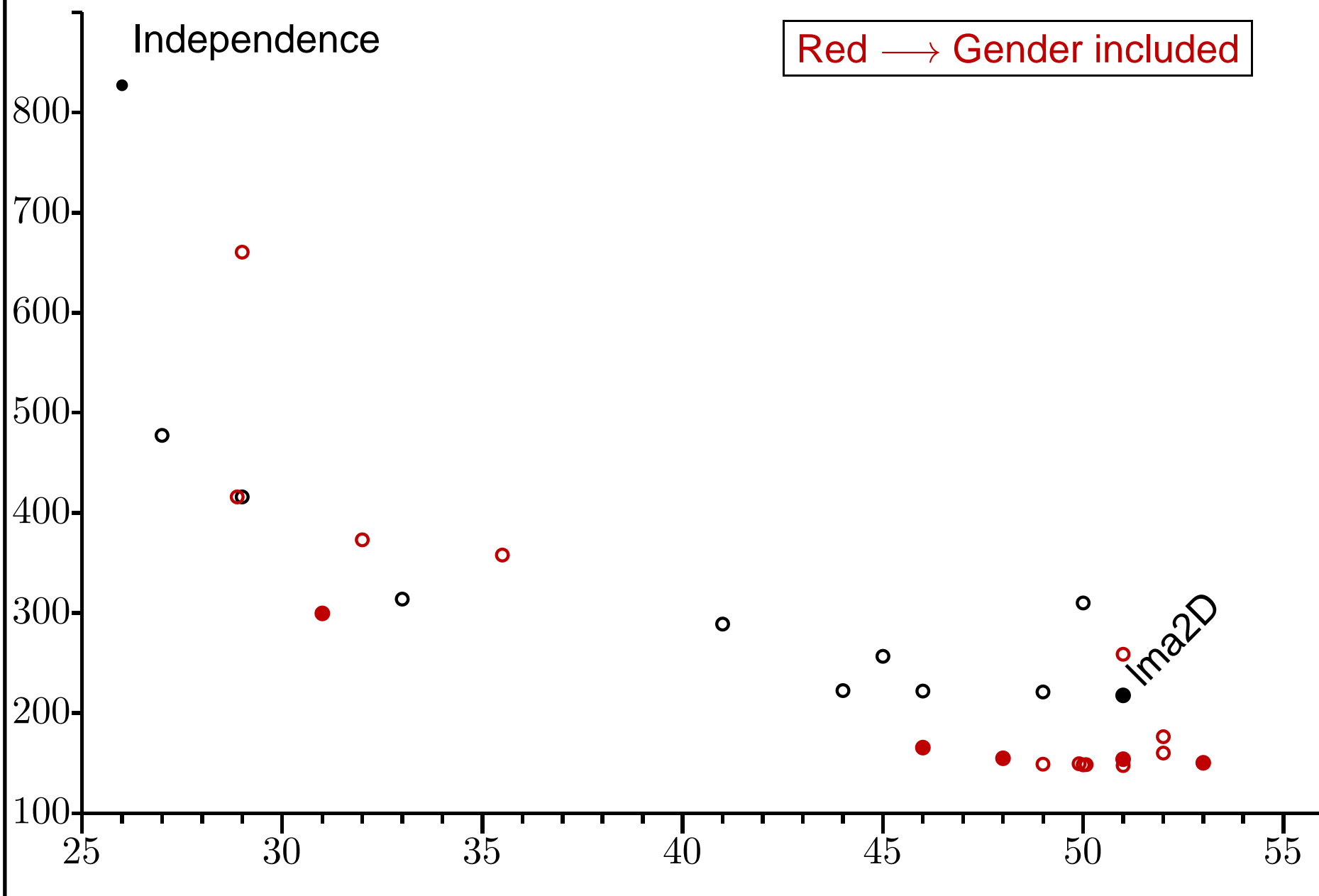
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-2lnlike versus Number Parameters w/ Gender

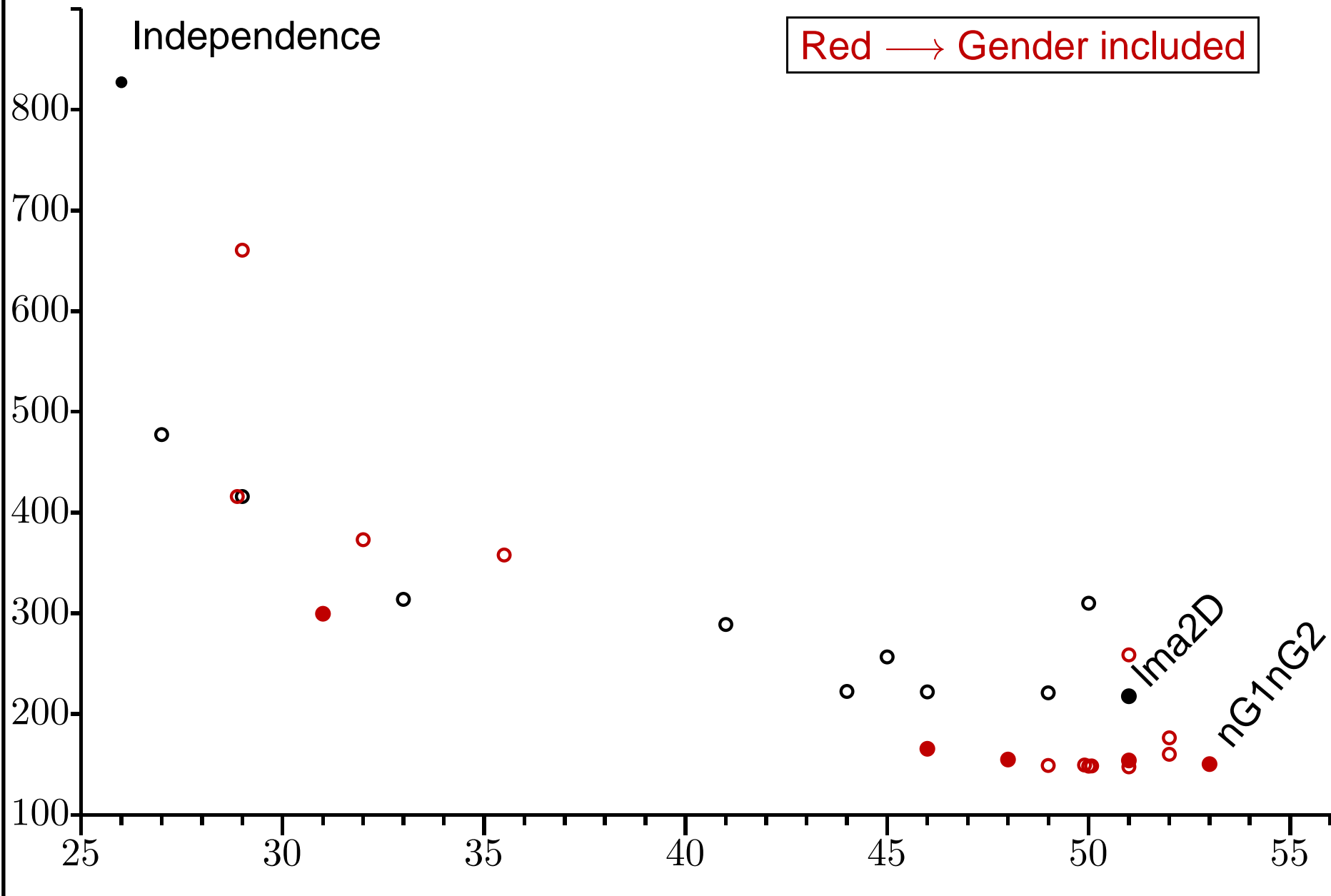
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-2lnlike versus Number Parameters w/ Gender

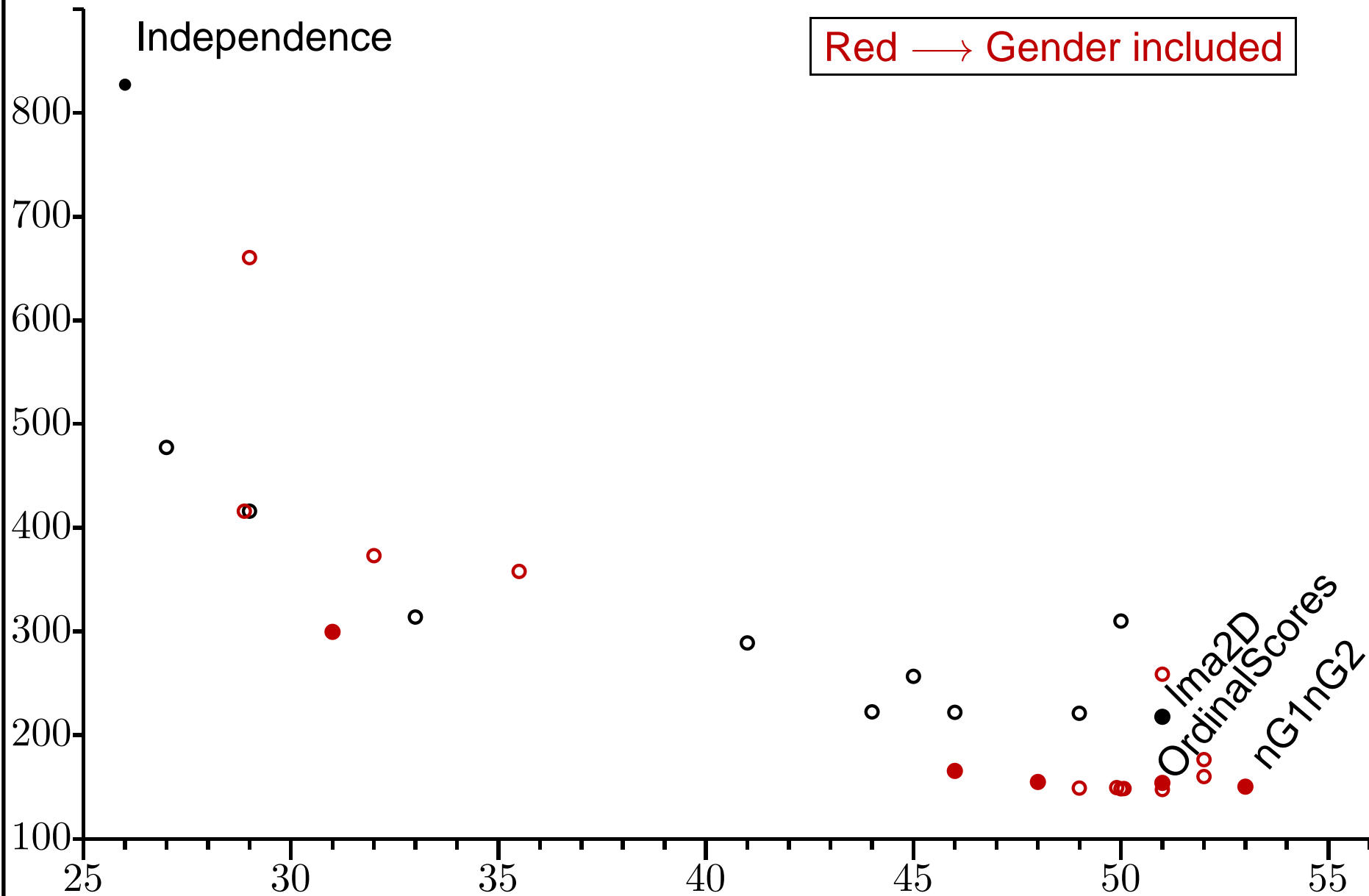
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-2lnlike versus Number Parameters w/ Gender

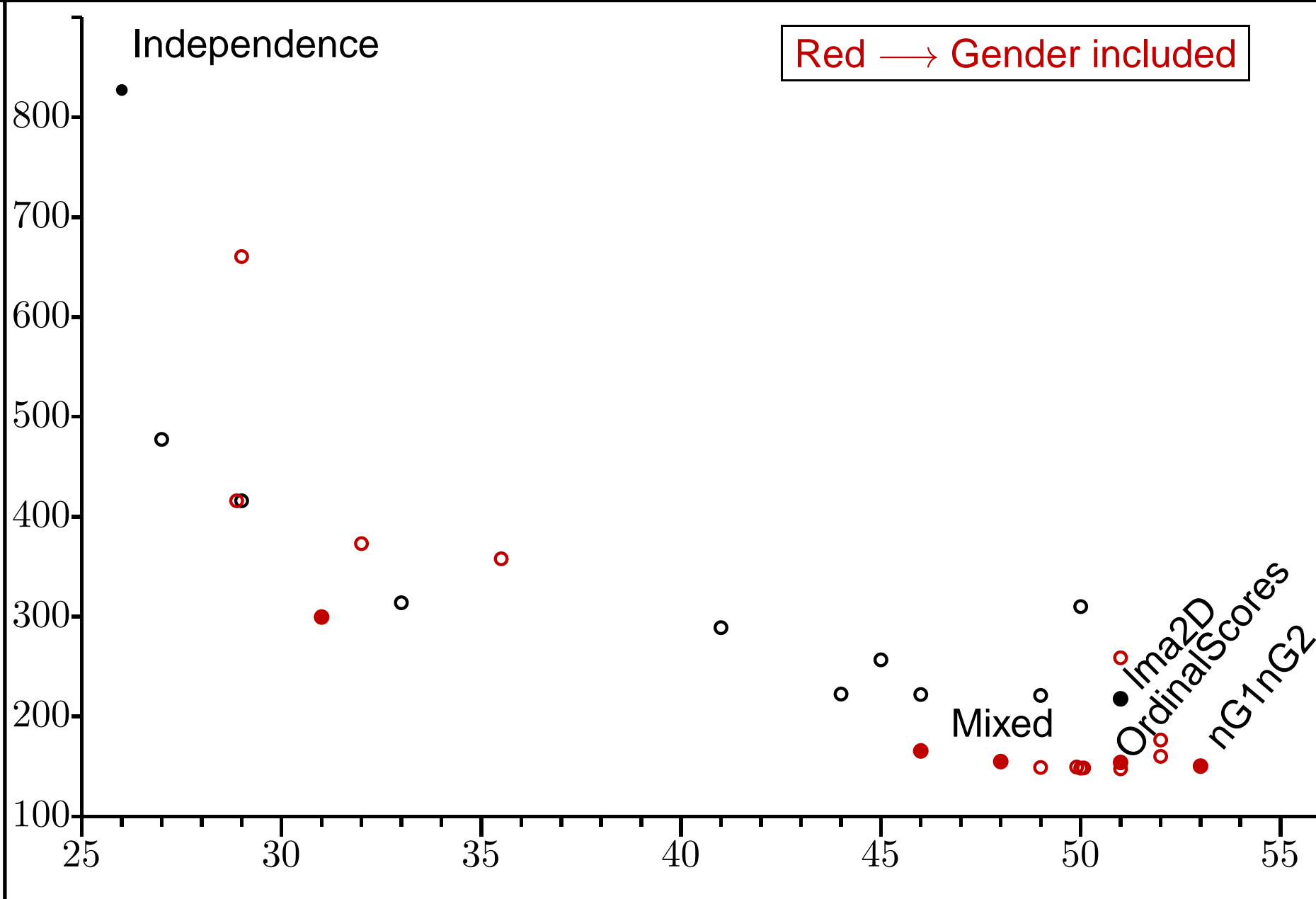
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-2lnlike versus Number Parameters w/ Gender

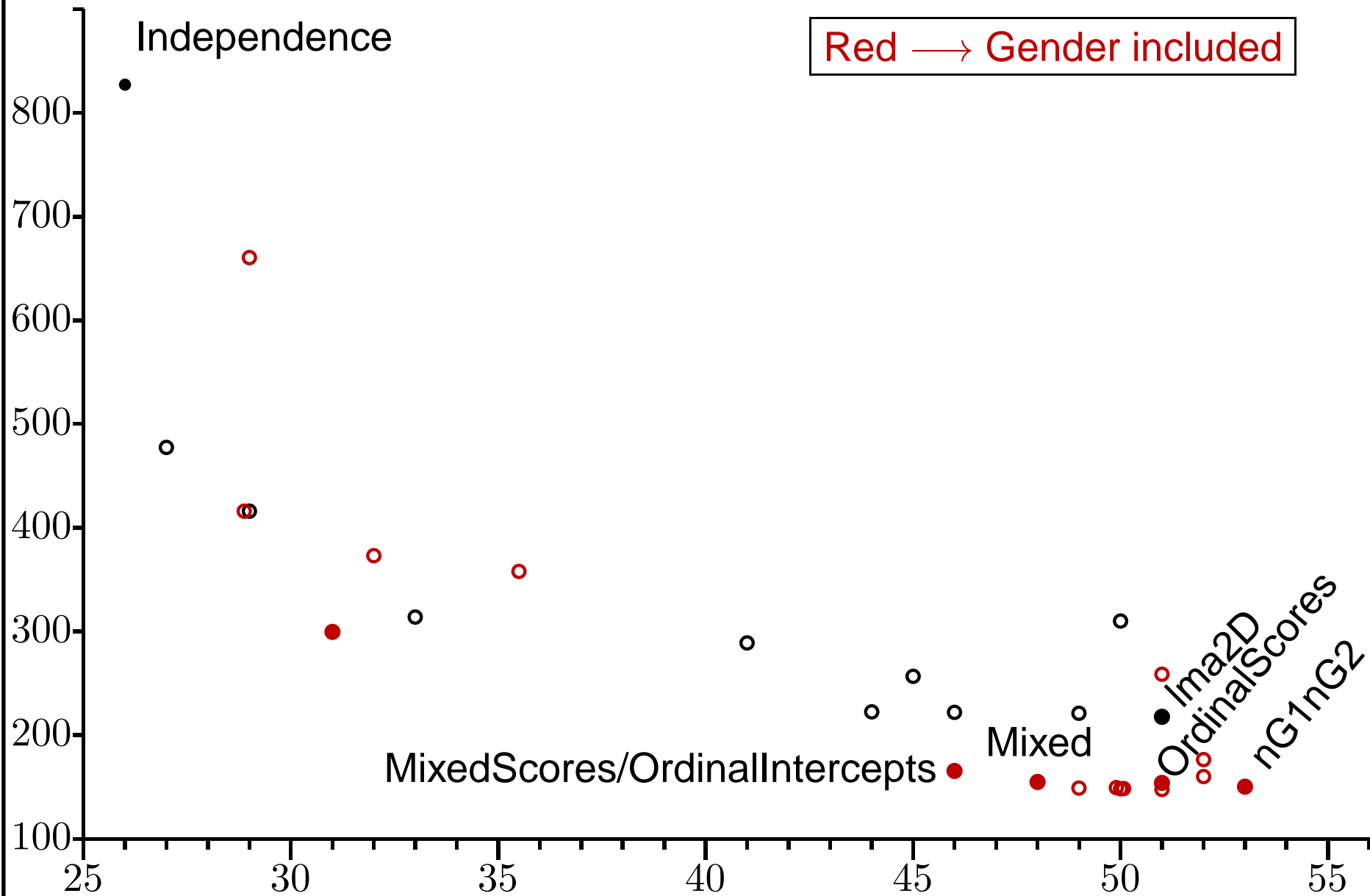
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-2lnlike versus Number Parameters w/ Gender

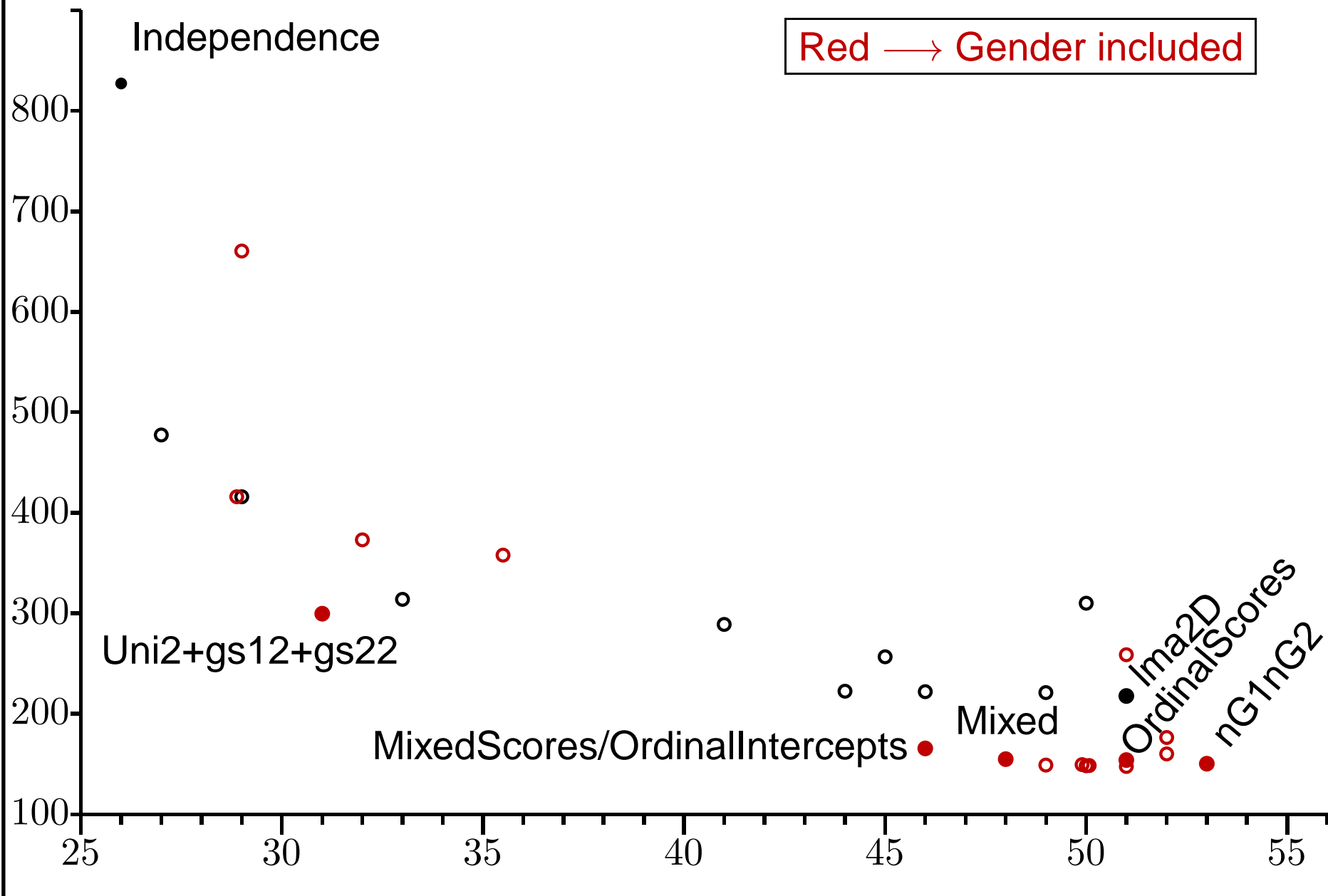
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Estimated Item Scale Values

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Parameters

● -2lnlike versus Number

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● **Estimated Item Scale Values**

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● Item Cumulative Probability

Curves

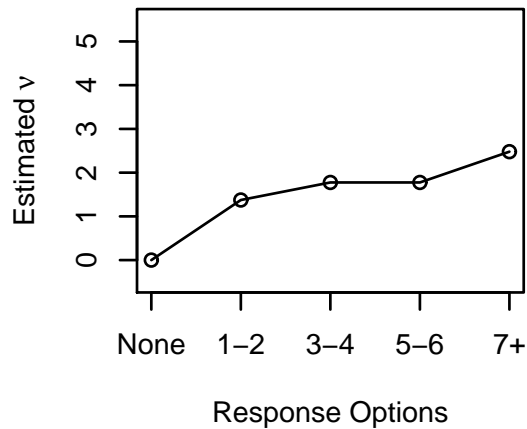
● Estimated Scale Values for

Gender

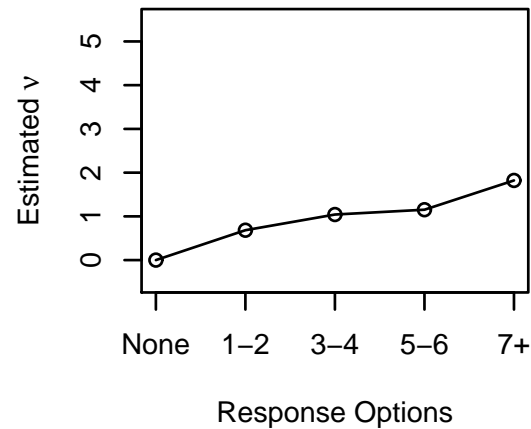
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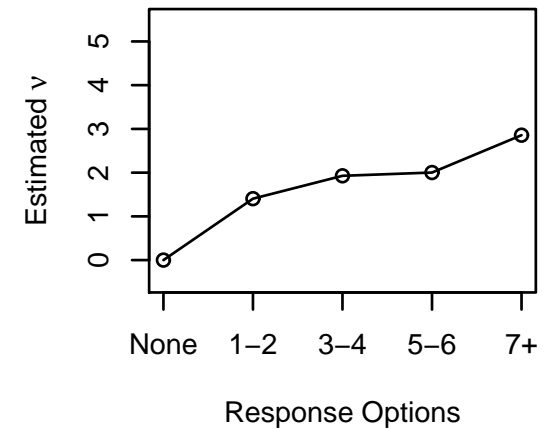
Got in Fight



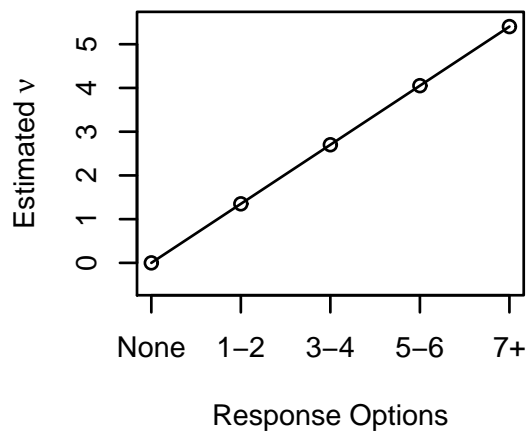
Threatened to hit/hurt



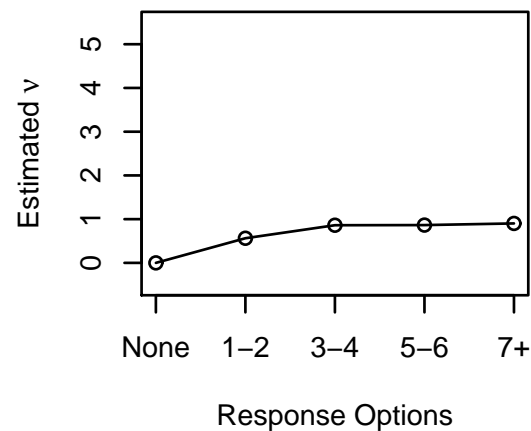
Hit back



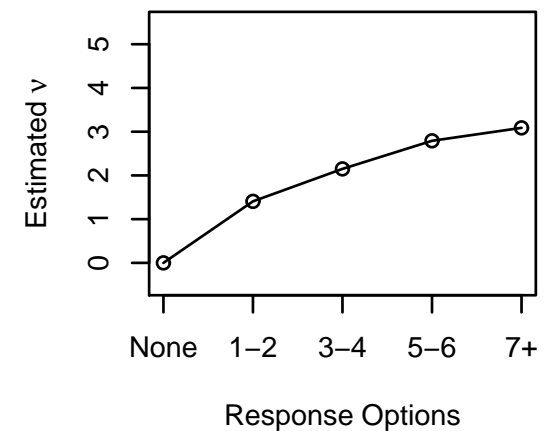
Upset others for fun



Help harass



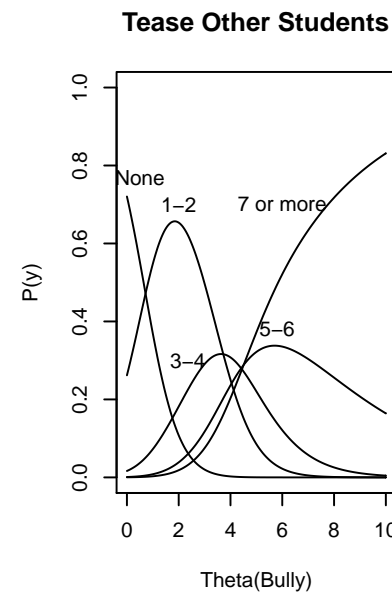
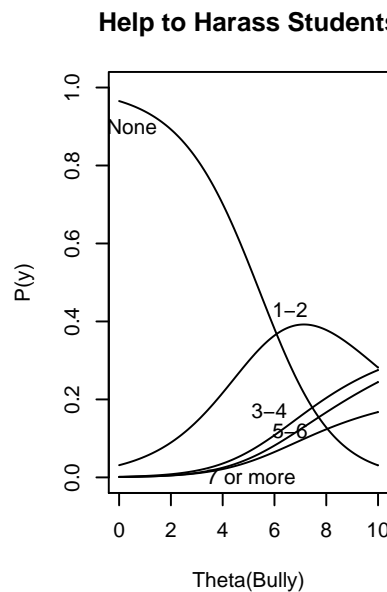
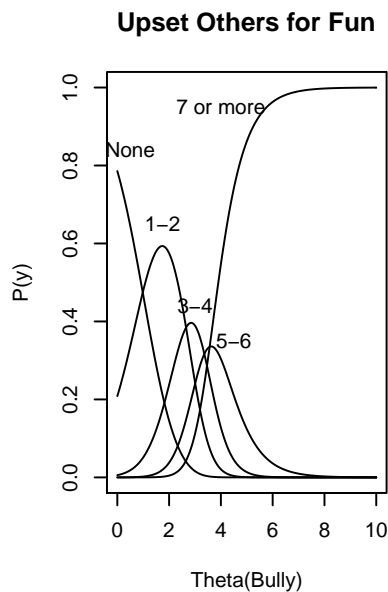
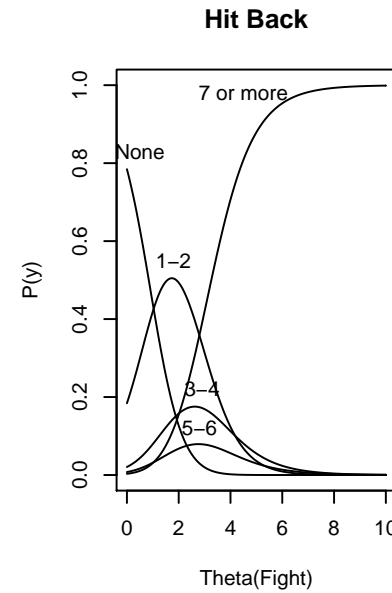
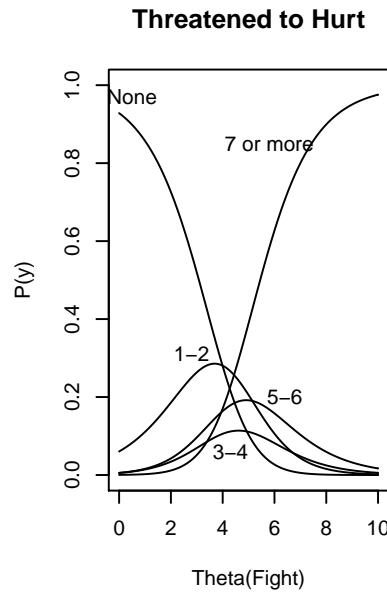
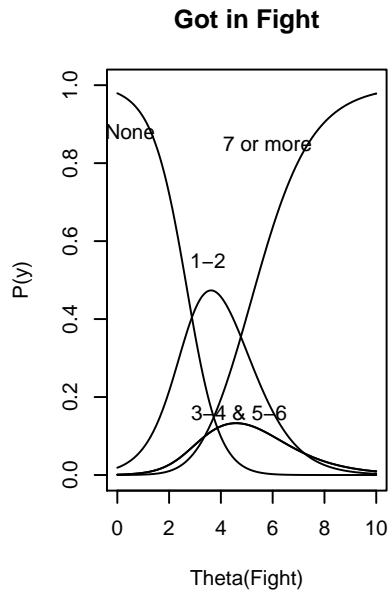
Tease Other Students





Item Characteristic Curves

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Item Cumulative Probability Curves

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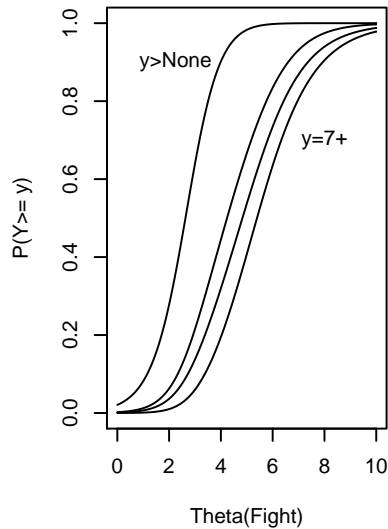
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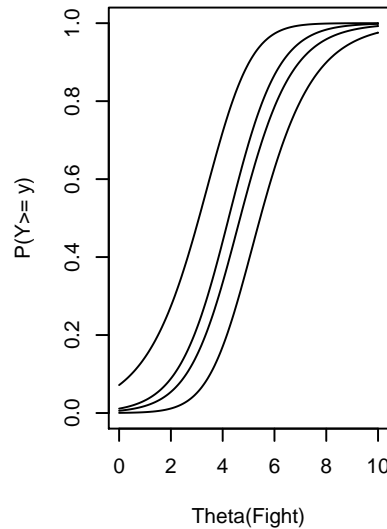
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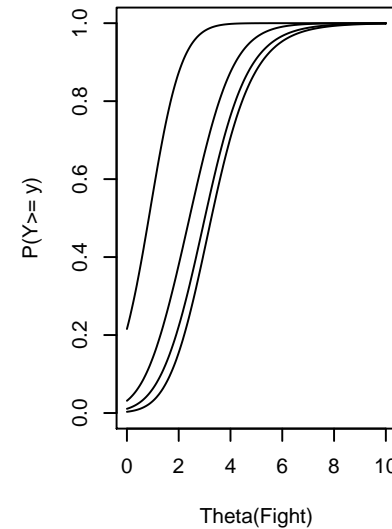
Got in Fight



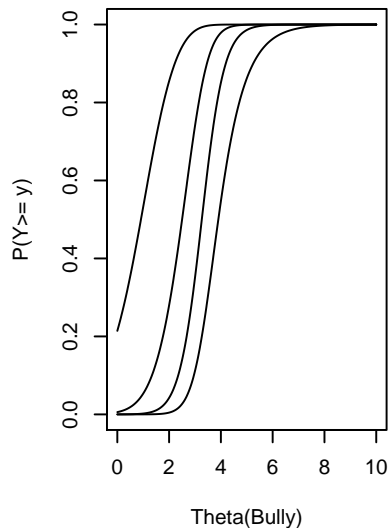
Threatened to Hurt



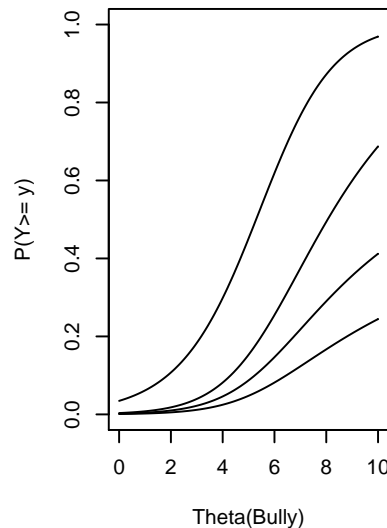
Hit Back



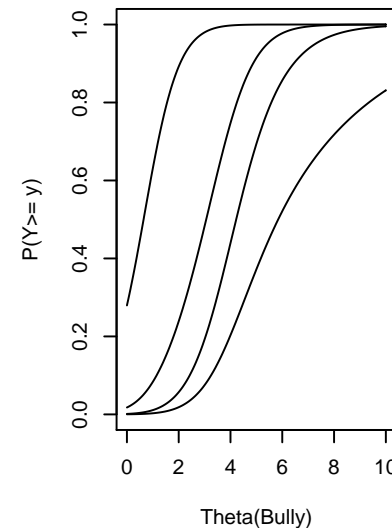
Upset Others for Fun



Help to Harass Students



Tease Other Students

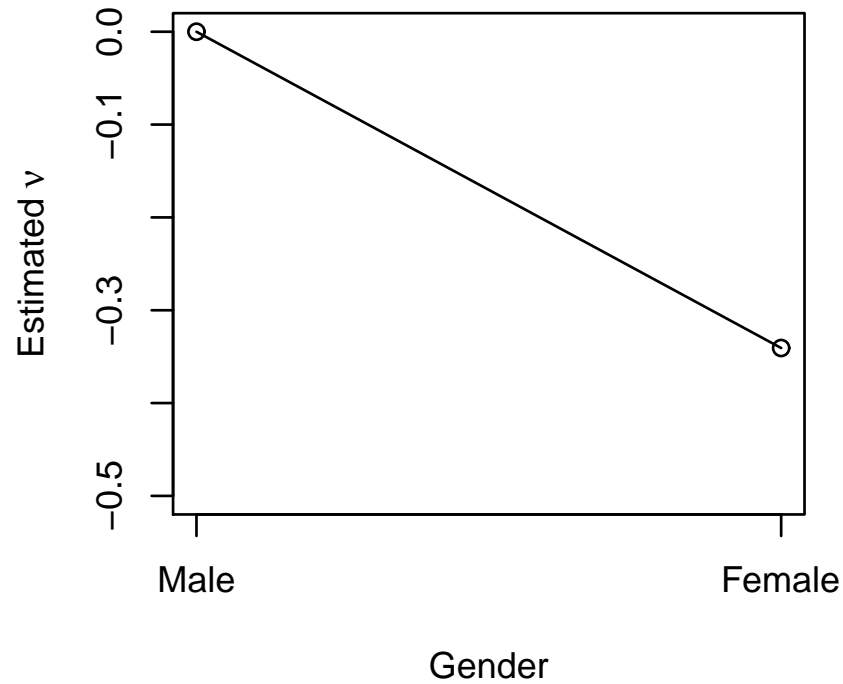




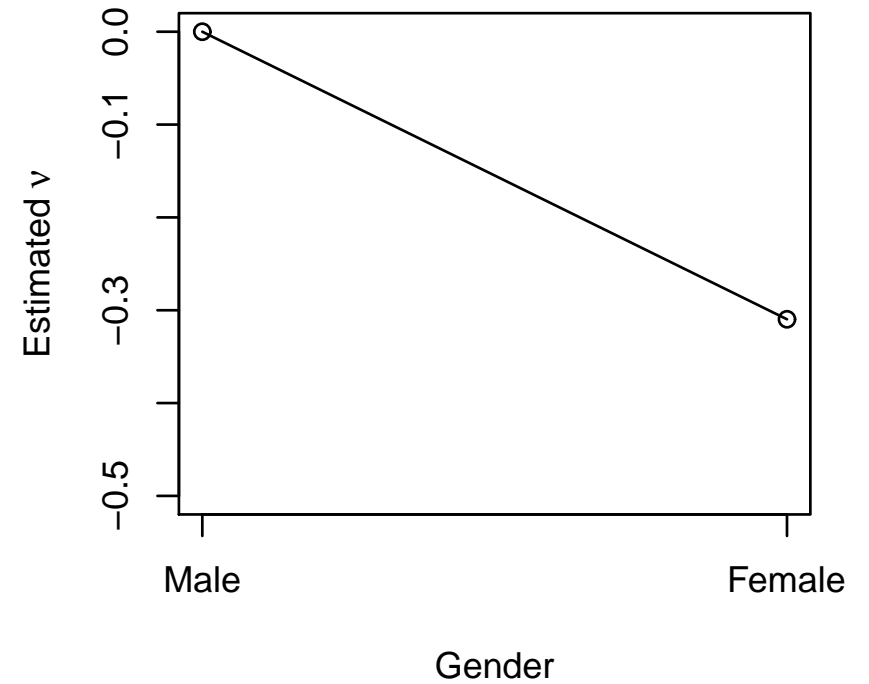
Estimated Scale Values for Gender

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Gender on Fight



Gender on Bully





Example Illustrated...

Noteworthy in today's example for multcategory items:

- Marginal distribution of traits are very skewed.

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Example Illustrated...

Noteworthy in today's example for multcategory items:

- Marginal distribution of traits are very skewed.
- Ordinal restrictions on responses options (and linear transformations).

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Example Illustrated...

Noteworthy in today's example for multcategory items:

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- Ordinal restrictions on responses options (and linear transformations).
- Ordinal restrictions on intercepts (“difficulties” or location parameters).

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Example Illustrated...

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- Marginal distribution of traits are very skewed.
- Ordinal restrictions on responses options (and linear transformations).
- Ordinal restrictions on intercepts (“difficulties” or location parameters).
- Covariate came in from the “right” (i.e., helps model underlying latent variable).

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Example Illustrated...

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- Ordinal restrictions on responses options (and linear transformations).
- Ordinal restrictions on intercepts (“difficulties” or location parameters).
- Covariate came in from the “right” (i.e., helps model underlying latent variable).
- The scale values for categories of the 3 three bully items are nearly identical from a uni-dimensional LMA model fit to all 9 bully items (and different estimation methods). Correlations $> .99$.

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Noteworthy in today's example for multcategory items:

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- Ordinal restrictions on intercepts (“difficulties” or location parameters).
- Covariate came in from the “right” (i.e., helps model underlying latent variable).
- The scale values for categories of the 3 three bully items are nearly identical from a uni-dimensional LMA model fit to all 9 bully items (and different estimation methods). Correlations $> .99$.

The Importance of this: Illustrates that major criticisms of conditional models do not hold up for LMA models as latent variable models:

- ◆ Models parameters are essentially the same.
- ◆ No interpretational difficulty.



Estimation Developments

The major problem is the size of a table (i.e., number of items/categories), but not the number of latent variables.

- The largest problem that I've successfully fit using ℓ_{EM} (Vermunt, 1997) is $2^{12} = 4096$ response patterns.

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- Models in the Rasch family can be fit by **pseudo-likelihood estimation** in any program that can fit conditional logistic regression models (Anderson, Li & Vermunt, 2007) and can include covariates.

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- For models with estimated category scores, an **experimental algorithm**.

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Experimental Algorithm

For models with estimated category scores, an experimental algorithm that iteratively fits conditional logistic regressions using MLE to estimate scale values and a pseudo-likelihood step to estimate the association parameters.

- Takes advantage of conditional specification of models.

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- Converges relatively quickly.

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- Estimation.
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- Further comparisons with traditional IRT methods (i.e., estimation of item parameters and individuals' values on latent variables).



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- Applications, especially those with multiple latent variables (e.g., testlets, Q-matrix).
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- Estimation.
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- Multidimensional, partially-compensatory models → leads to higher-way interactions in model for the data where the higher-way interactions have higher-way decompositions (Tucker 3-mode and other higher-way type decompositions).



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Will be able to download SAS/NLP programs used in this talk and various papers from

http://faculty.ed.uiuc.edu/cja/homepage/software_index.html

and slides from

<http://faculty.ed.uiuc.edu/cja/homepage>



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A: Who has the final responsibility to decide if a law is constitutional or not?

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American National Elections Pilot (1998)

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Anderson, C.J., Verkuilen, J.V., & Peyton, B.L. (in press). *Journal of Educational and Behavioral Statistics*.



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- Ordering of response the options unknown.



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- Ordering of response the options unknown.
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- Dealing with “Don’t Know”.
- Different number of response options.



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- Ordering of response the options unknown.
- Dealing with “Don’t Know”.
- Different number of response options.
- Scoring of responses needed.



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- Ordering of response the options unknown.
- Dealing with “Don’t Know”.
- Different number of response options.
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- Latent variable structure unknown.



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 - ◆ One dominant underlying dimension of “Knowledge.”



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 - ◆ One dominant underlying dimension of “Knowledge.”
 - ◆ Two correlated latent variables: Structure of political system (items A & B) and Party in Power (items C & D)



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- How do the instructions given to respondents affect all of this?



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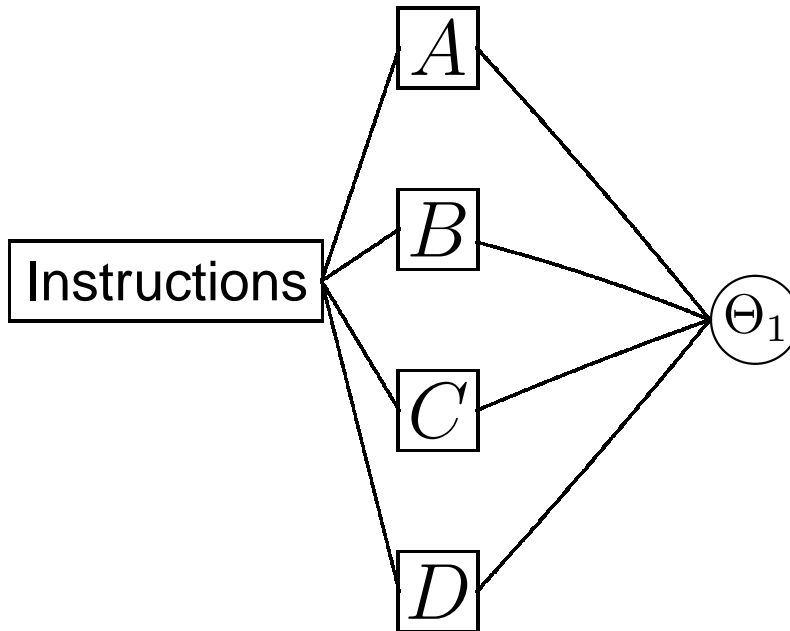
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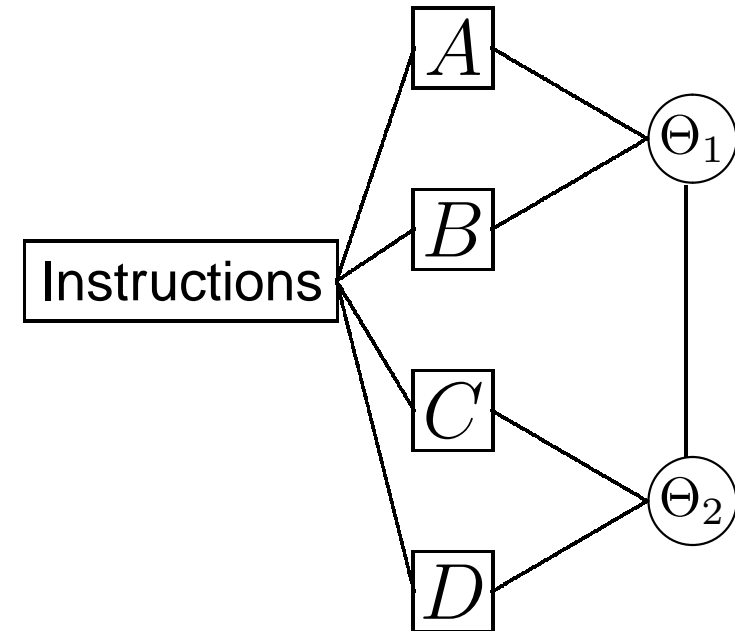
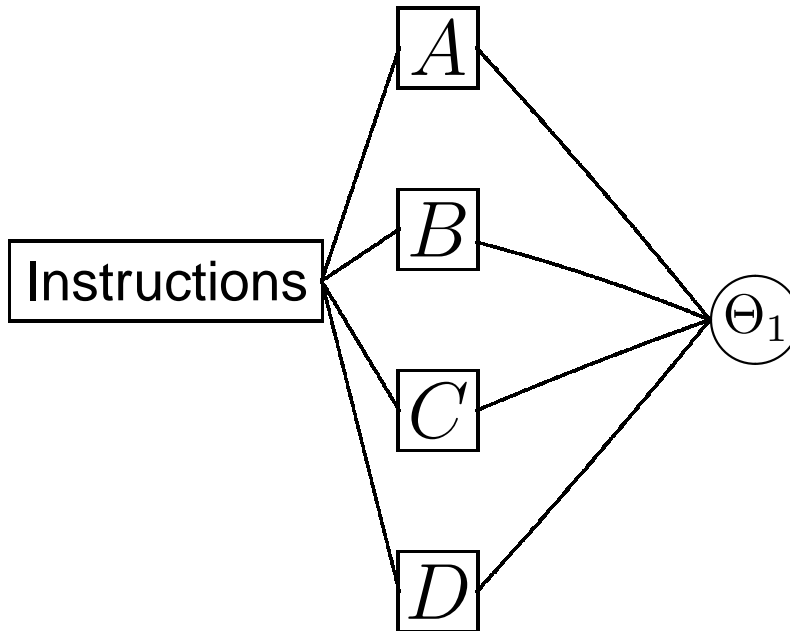
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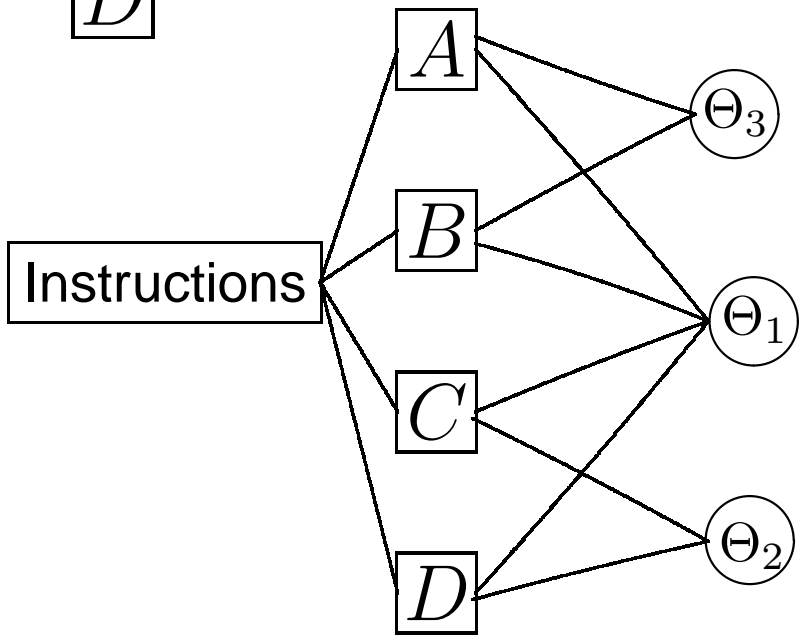
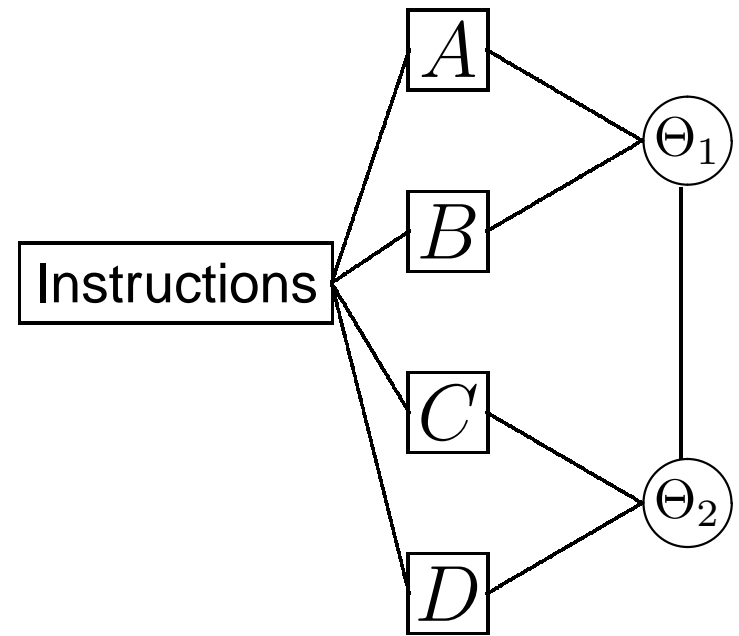
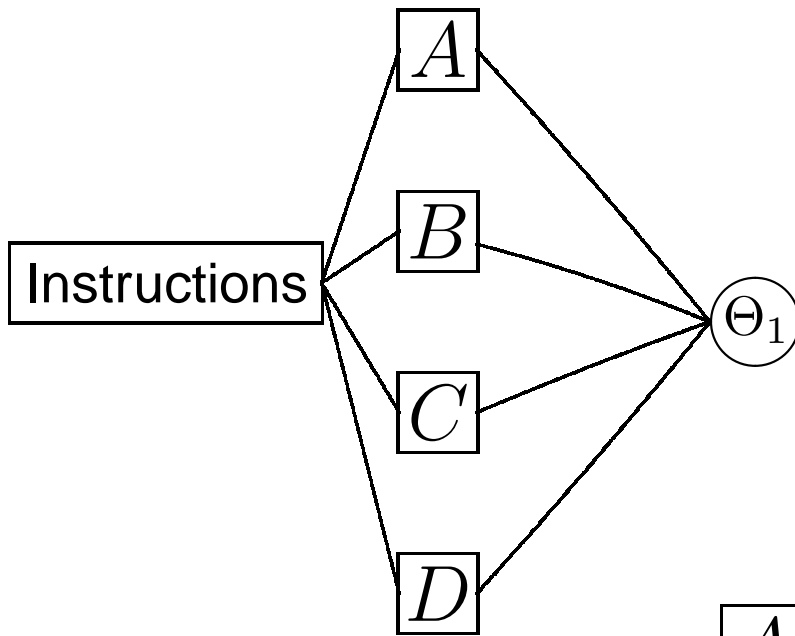




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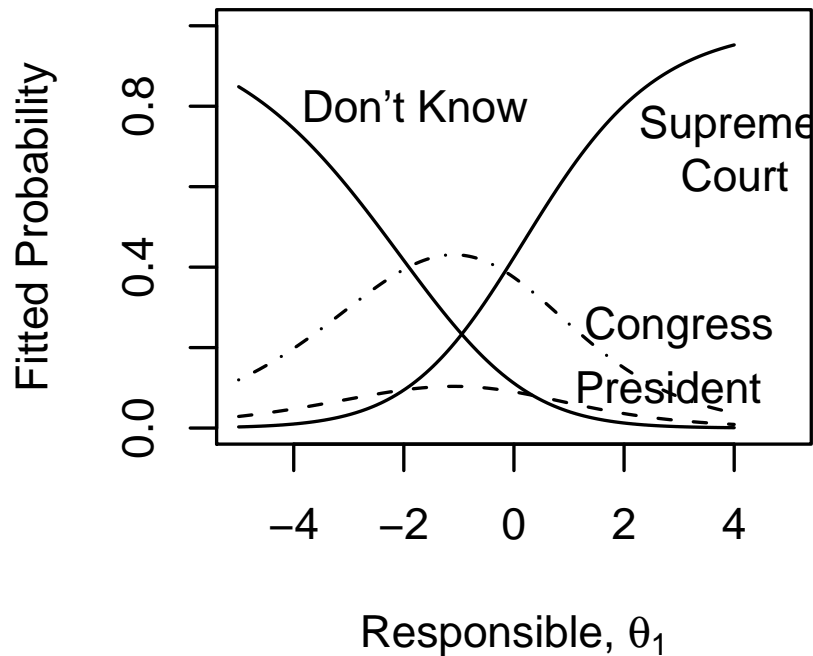
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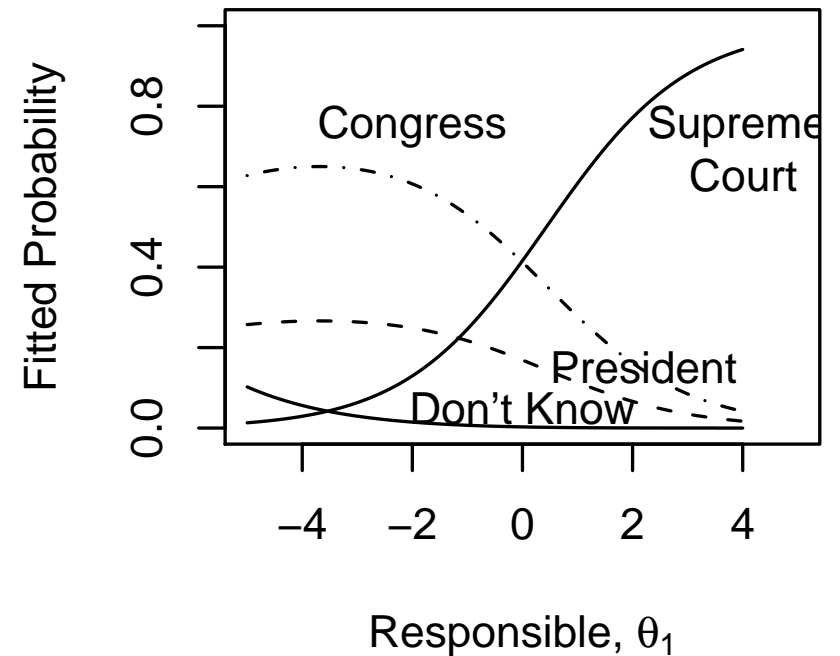
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Standard Instructions



Encourage Guessing



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